Tutorial

Characterizing the Generalization Error of Machine Learning Algorithms via Information Measures

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Preliminaries

Information Measures

- KL divergence: $\operatorname{KL}(P \| Q) \triangleq \int_{\mathcal{X}} \log\left(\frac{dP}{dQ}\right) dP$
- Symmetrized KL divergence (Jeffrey's divergence)

 $D_{\mathrm{SKL}}(P||Q) \triangleq \mathrm{KL}(P||Q) + \mathrm{KL}(Q||P).$

- Mutual information: $I(X; Y) \triangleq KL(P_{X,Y} || P_X \otimes P_Y)$
- ► Lautum information [Palomar and Verdú, 2008]: $L(X; Y) \triangleq KL(P_X \otimes P_Y || P_{X,Y})$
- ► Symmetrized KL information [Aminian et al., 2015]:

$$I_{\rm SKL}(X;Y) \triangleq D_{\rm SKL}(P_{X,Y} || P_X \otimes P_Y) = I(X;Y) + L(X;Y)$$

Information-theoretic Generalization Bounds

Lemma ([Xu and Raginsky, 2017])

Suppose $\ell(w, Z)$ is σ -sub-Gaussian under $Z \sim \mu$ for all $w \in W$, then

$$|\operatorname{gen}(\mu, P_{W|S})| \leq \sqrt{\frac{2\sigma^2}{n}}\operatorname{I}(S; W)$$

- > Depends on every ingredient in the supervised learning problem
- \blacktriangleright Reducing dependence between W and S leads to better generalization bound
- This bound is only tight if I(S; W) = 0 and $gen(\mu, P_{W|S}) = 0$
- ► Multiple techniques to improve this result, including ISMI [Bu et al., 2020], CMI [Steinke and Zakynthinou, 2020], *f*-CMI [Harutyunyan et al., 2021], △*L*-CMI [Wang and Mao, 2023]

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Regularized ERM problem

- ▶ How can we use this result to develop a better learning algorithm?
- Regularizing mutual information I(S; W) during ERM

$$P_{W|S}^{\star} = \underset{P_{W|S}}{\operatorname{arg\,min}} \left(\mathbb{E}_{P_{W,S}}[L_{E}(W,S)] + \frac{1}{\gamma} \mathbf{I}(S;W) \right)$$

- \blacktriangleright inverse temperature $\gamma \geq \mathbf{0}$ balances between fitting and generalization
- ▶ Replacing I(*S*; *W*) with $KL(P_{W|S} || \pi(W) | P_S)$ for any prior $\pi(W)$
- ▶ It gives information risk minimization (IRM) problem

$$P_{W|S}^{\star} = \arg\min_{P_{W|S}} \left(\mathbb{E}_{P_{W,S}} [L_E(W, S)] + \frac{1}{\gamma} \mathrm{KL}(P_{W|S} || \pi(W) | P_S) \right)$$

Information Risk Minimization

Lemma ([Zhang, 2006, Xu and Raginsky, 2017])

Solution to IRM problem is $(\gamma, \pi(w), L_E(w, s))$ -Gibbs distribution

$$P_{W|S}^{\gamma}(w|s) riangleq rac{\pi(w)e^{-\gamma L_E(w,s)}}{V(s,\gamma)}, \quad \gamma \geq 0,$$

where $V(s,\gamma) \triangleq \int \pi(w) e^{-\gamma L_E(w,s)} dw$ is partition function.

Proof.

For any learning algorithm $P_{W|S}$ with fixed S = s,

$$0 \leq \operatorname{KL}(P_{W|S=s} \| P_{W|S=s}^{\gamma})$$

$$= \mathbb{E}_{P_{W|S=s}} \left[\log \frac{P_{W|S=s} \cdot V(s,\gamma)}{\pi(W) \cdot e^{-\gamma L_{E}(w,s)}} \right]$$

$$= \operatorname{KL}(P_{W|S=s} \| \pi(W)) + \log V(s,\gamma) + \gamma \mathbb{E}_{P_{W|S=s}} [L_{E}(w,s)].$$

$$\min_{W|S} \mathbb{E}_{P_{W|S=s}} \left[L_{E}(W,s) \right] + \frac{1}{\gamma} \operatorname{KL}(P_{W|S=s} \| \pi) = -\frac{1}{\gamma} \log V(s,\gamma).$$

Gibbs Algorithm

We focus on the generalization error of Gibbs algorithm (distribution)

 $(\gamma, \pi(w), L_E(w, s))$ -Gibbs distribution:

$${\sf P}^{\gamma}_{W|S}(w|s) riangleq rac{\pi(w)e^{-\gamma L_E(W,s)}}{V(s,\gamma)}, \hspace{1em} \gamma \geq 0$$

where

- \blacktriangleright inverse temperature $\gamma,$ reduces to standard ERM if $\gamma \rightarrow \infty$
- $\pi(w)$ arbitrary prior distribution of W
- $V(s,\gamma) \triangleq \int \pi(w) e^{-\gamma L_E(w,s)} dw$ partition function

Practical Implementation of Gibbs algorithm

- ► Stochastic Gradient Langevin Dynamics (SGLD)
- Metropolis adjusted Langevin algorithm (MALA)

The SGLD can be viewed as the noisy version of SGD,

$$W_{k+1} = W_k - \eta_t
abla L_E(W_k, s) + \sqrt{rac{2\eta_t}{\gamma}} \zeta_k, \quad k = 0, 1, \cdots,$$

where ζ_k standard Gaussian random vector; $\eta_t > 0$ step size.

- ► [Raginsky et al., 2017] shows that $P_{W_k|S}$ induced by SGLD converges to $(\gamma, \pi(W_0), L_E(w_k, s))$ -Gibbs distribution for sufficiently large k
- ▶ MALA is SGLD with Metropolis rejection, faster convergence

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Expected Generalization Error

An exact characterization of generalization error for Gibbs algorithm

Theorem

For $(\gamma, \pi(w), L_E(w, s))$ -Gibbs algorithm,

$${\mathcal P}^{\gamma}_{W|S}(w|s)=rac{\pi(w)e^{-\gamma L_E(w,s)}}{V(s,\gamma)}, \quad \gamma>0,$$

the expected generalization error is

$$\overline{\operatorname{gen}}(P_{W|S}^{\gamma}, P_S) = \frac{\operatorname{I}_{\operatorname{SKL}}(W; S)}{\gamma}.$$

- ▶ Highlights the fundamental role of $I_{SKL}(W; S)$ in learning theory
- Holds even for non-i.i.d training samples

G. Aminian*, Y. Bu*, L. Toni, M. R. Rodrigues, G. W. Wornell. "An Exact Characterization of the Generalization Error for the Gibbs Algorithm," in *Proc. Conference on Neural Information Processing Systems* (NeurIPS), Dec. 2021.

Generalization Error of Gibbs Algorithm

Theorem

For Gibbs algorithm
$$P_{W|S}^{\gamma}(w|s) = \frac{\pi(w)e^{-\gamma L_E(w,s)}}{V(s,\gamma)}$$
,
 $\overline{\operatorname{gen}}(P_{W|S}^{\gamma}, P_S) = \frac{\operatorname{I}_{\mathrm{SKL}}(W; S)}{\gamma}$.

Sketch of Proof:

Symmetrized KL information can be written as $I_{SKL}(W; S) = \mathbb{E}_{P_{W,S}}[\log(\frac{P_{W|S}^{\gamma}}{P_{W}})] + \mathbb{E}_{P_{W} \otimes P_{S}}[\log(\frac{P_{W}}{P_{W|S}^{\gamma}})]$ $= \mathbb{E}_{P_{W,S}}[\log(P_{W|S}^{\gamma})] - \mathbb{E}_{P_{W} \otimes P_{S}}[\log(P_{W|S}^{\gamma})]$

Note that $P_{W,S}$ and $P_W \otimes P_S$ share the same marginal distribution,

$$\begin{split} \mathrm{I}_{\mathrm{SKL}}(W;S) &= \mathbb{E}_{P_{W,S}}[-\gamma L_{\mathcal{E}}(W,S)] - \mathbb{E}_{P_{W}\otimes P_{S}}[-\gamma L_{\mathcal{E}}(W,S)] \\ &= \gamma \overline{\mathrm{gen}}(P_{W|S}^{\gamma},P_{S}) \end{split}$$

Theorem

log $V(s, \gamma)$ is convex and differentiable infinitely many times with respect to γ . In particular,

$$\mathbb{E}_{\gamma}[L_{E}(W,s)] = -\frac{\partial \log V(s,\gamma)}{\partial \gamma},$$

$$Var_{\gamma}[L_{E}(W,s)] = \frac{\partial^{2} \log V(s,\gamma)}{\partial \gamma^{2}},$$

where $\mathbb{E}_{\gamma}[\cdot] \triangleq \mathbb{E}_{P_{W|S=s}^{\gamma}}[\cdot]$, and $\operatorname{Var}_{\gamma}[L_{E}(W, s)] \triangleq \mathbb{E}_{\gamma}[L_{E}(W, s)^{2}] - \mathbb{E}_{\gamma}[L_{E}(W, s)]^{2}$.

Expected empirical risk of the Gibbs algorithm is non-increasing w.r.t γ

- Monoticity: $L_E(W, s)$ is non-increasing with γ
- ▶ Sub-Gaussianity: $L_E(W, s)$ is sub-Gaussian under Gibbs algorithm if $\operatorname{Var}_{\gamma}[L_E(W, s)]$ is bounded

Perlaza, Samir M., Gaetan Bisson, Iñaki Esnaola, Alain Jean-Marie, and Stefano Rini. "Empirical risk minimization with relative entropy regularization," *IEEE Trans. Inf. Theory*, 2024.

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Tighter Generalization Error Upper Bounds

Why do we care about upper bounds when we have exact characterization?

- Quantify how $\overline{\text{gen}}(P_{W|S}^{\gamma}, P_S)$ depends on number of i.i.d. samples n
- Useful when directly evaluating $I_{SKL}(W; S)$ is hard

Theorem

Suppose that

- $S = \{Z_i\}_{i=1}^n$ are *i.i.d* generated from the distribution P_Z
- $\ell(w, Z)$ is σ -sub-Gaussian
- $C_E \leq \frac{L(W;S)}{I(W;S)}$ for some $C_E \geq 0$,

$$\overline{\operatorname{gen}}({\mathcal{P}}_{W|{\mathcal{S}}}^{\gamma},{\mathcal{P}}_{{\mathcal{S}}})\leq rac{2\sigma^2\gamma}{(1+C_{{\mathcal{E}}})n}.$$

G. Aminian*, Y. Bu*, L. Toni, M. R. Rodrigues, G. W. Wornell. "An Exact Characterization of the Generalization Error for the Gibbs Algorithm," in *Proc. Conference on Neural Information Processing Systems* (NeurIPS), Dec. 2021.

Tighter Generalization Error Upper Bounds

Sketch of Proof:

Recall the mutual information-based bound,

$$egin{aligned} &\sqrt{rac{2\sigma^2}{n}}\mathrm{I}(S;W) \geq \overline{\mathrm{gen}}(P_{W|S}^\gamma,P_S) = rac{\mathrm{I}(W;S)+L(W;S)}{\gamma} \ &\geq rac{(1+C_E)}{\gamma}\mathrm{I}(W;S) \ &\overline{\mathrm{gen}}(P_{W|S}^\gamma,P_S) \leq \sqrt{rac{2\sigma^2}{n}}\mathrm{I}(S;W) \leq rac{2\sigma^2\gamma}{(1+C_E)n} \end{aligned}$$

[Choice of C_E]

- $C_E = 0$ is always valid, which gives $\overline{\text{gen}}(P_{W|S}^{\gamma}, P_S) \leq \frac{2\sigma^2 \gamma}{n}$
- ▶ $C_E = 1$, $L(S; W) \ge I(S; W)$ holds for any Gaussian channel $P_{W|S}$

Example: Mean Estimation

- ▶ Learning mean $\mu \in \mathbb{R}^d$ of Z using n i.i.d training samples $S = \{z_i\}_{i=1}^n$
- ▶ Not necessary Gaussian, but covariance matrix $\Sigma_Z = \sigma_Z^2 I_d$
- Mean-squared loss $\ell(w, z) = ||z w||_2^2$
- Gaussian prior $\pi(w) = \mathcal{N}(\mu_0, \sigma_0^2 I_d)$
- ► Then, $(\gamma, \mathcal{N}(\mu_0, \sigma_0^2 I_d), L_E(w, s))$ -Gibbs algorithm is given by the following Gaussian posterior

$$P_{W|S}^{\gamma}(\boldsymbol{w}|\boldsymbol{z}^{n}) \sim \mathcal{N}\Big(\alpha \boldsymbol{\mu}_{0} + (1-\alpha) \bar{\boldsymbol{z}}, \alpha \sigma_{0}^{2} \boldsymbol{I}_{d}\Big),$$

with

$$\alpha \triangleq rac{1}{2\sigma_0^2\gamma + 1}, \quad ar{z} \triangleq rac{1}{n}\sum_{i=1}^n z_i.$$

Example: Mean Estimation

Since $P_{W|S}^{\gamma}$ is Gaussian,

$$\begin{split} \mathrm{I}(S;W) &= \frac{d\sigma_0^2 \sigma_Z^2 \gamma}{(n\sigma_0^2 + \frac{1}{2\gamma})} - \mathrm{KL}(P_W \| \mathcal{N}(\boldsymbol{\mu}_W, \sigma_1^2 \boldsymbol{I}_d)), \\ \mathcal{L}(S;W) &= \frac{d\sigma_0^2 \sigma_Z^2 \gamma}{(n\sigma_0^2 + \frac{1}{2\gamma})} + \mathrm{KL}(P_W \| \mathcal{N}(\boldsymbol{\mu}_W, \sigma_1^2 \boldsymbol{I}_d)), \end{split}$$

with $\mu_W = \alpha \mu_0 + (1 - \alpha) \mu_.$

The generalization error can be computed exactly as:

$$\overline{\operatorname{gen}}(\boldsymbol{P}_{W|S}^{\gamma},\boldsymbol{P}_{S}) = \frac{\operatorname{I}_{\operatorname{SKL}}(W;S)}{\gamma} = \frac{2d\sigma_{0}^{2}\sigma_{Z}^{2}}{n(\sigma_{0}^{2} + \frac{1}{2\gamma})}.$$

As a comparison, the ISMI-based bound gives a sub-optimal bound $\mathcal{O}\left(1/\sqrt{n}\right)$, as $n \to \infty$.

Check Point

Generalization error or empirical risk is one part of the story

Our goal is to design (or guide the design) algorithms that minimize population risk.

There are three elements in $(\gamma, \pi(w), L_E(w, s))$ -Gibbs algorithm

- \blacktriangleright inverse temperature $\gamma \longrightarrow \operatorname{Optimal}$ hyper-parameter
- empirical risk $L_E(w, s)$, or model family \longrightarrow Information criteria for model selection
- prior distribution $\pi(w) \longrightarrow$ Transfer learning

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Expected Test Loss

For fixed training data s and testing data s', consider expected test loss

 $L_P(\gamma, s, s') \triangleq \mathbb{E}_{\gamma}[L_E(W, s')],$

and expected generalization error

 $\overline{\operatorname{gen}}(\gamma, \boldsymbol{s}, \boldsymbol{s}') \triangleq \mathbb{E}_{\gamma}[L_{E}(W, \boldsymbol{s}') - L_{E}(W, \boldsymbol{s})].$

Theorem

For $\gamma \geq 0$ such that $\log V(s,\gamma) < \infty$, the first order derivative of the expected test loss is given by

$$\frac{\partial}{\partial \gamma} L_{\mathcal{P}}(\gamma, s, s') = -\operatorname{Cov}_{\gamma}[L_{\mathcal{E}}(W, s'), L_{\mathcal{E}}(W, s)],$$

with

 $\operatorname{Cov}_{\gamma}[L_{E}(W,s'),L_{E}(W,s)] \triangleq \mathbb{E}_{\gamma}[L_{E}(W,s)L_{E}(W,s')] - \mathbb{E}_{\gamma}[L_{E}(W,s)]\mathbb{E}_{\gamma}[L_{E}(W,s')].$

 $\operatorname{Cov}_{\gamma}[L_{E}(W, s'), L_{E}(W, s)]$ can be positive/negative, no monotonicity

Y. Bu, "Towards Optimal Inverse Temperature in the Gibbs Algorithm," in IEEE ISIT 2024

Expected Generalization Error

Corollary

For $\gamma \ge 0$ such that $\log V(s, \gamma) < \infty$, the first order derivative of the expected generalization error is given by

$$\frac{\partial}{\partial \gamma} \overline{\operatorname{gen}}(\gamma, \boldsymbol{s}, \boldsymbol{s}') = \operatorname{Var}_{\gamma}(L_{E}(W, \boldsymbol{s})) - \operatorname{Cov}_{\gamma}[L_{E}(W, \boldsymbol{s}'), L_{E}(W, \boldsymbol{s})].$$

▶ Cannot show that the gen is non-decreasing, Cauchy-Schwarz Inequality only guarantees that

 $\left|\operatorname{Cov}_{\gamma}[L_{E}(W,s'),L_{E}(W,s)]\right| \leq \sqrt{\operatorname{Var}_{\gamma}(L_{E}(W,s))\operatorname{Var}_{\gamma}(L_{E}(W,s'))}.$

- [Aminian et al., 2021] provides a bound of order O(^γ/_n) by simply combining the I_{SKL} characterization with the MI bound, which may hint that gen is always increasing with γ.
- \blacktriangleright However, we will illustrate how gen rises from zero and then decreases as γ increases.

Example: Mean Estimation

 $(\gamma, \mathcal{N}(\mu_0, \sigma_0^2 I_d), L_E(w, s))$ -Gibbs algorithm is given by the following Gaussian posterior

$$P^{\gamma}_{W|S}(\boldsymbol{w}|\boldsymbol{z}^n) \sim \mathcal{N}\Big(\alpha \boldsymbol{\mu}_0 + (1-\alpha) \bar{\boldsymbol{z}}, \alpha \sigma_0^2 \boldsymbol{I}_d\Big)$$

Population risk has the following exact characterization

$$L_{P}(P_{W|S}^{\gamma}, P_{S}) = \underbrace{\frac{4d\sigma_{0}^{2}\sigma_{Z}^{2}\gamma}{n(1+2\sigma_{0}^{2}\gamma)}}_{\text{generalization error}} + \underbrace{\frac{\|\mu_{0} - \mu\|_{2}^{2} + d\sigma_{z}^{2}/n}{(1+2\sigma_{0}^{2}\gamma)^{2}} + \frac{d\sigma_{0}^{2}}{1+2\sigma_{0}^{2}\gamma} + \frac{n-1}{n}d\sigma_{Z}^{2}}_{\text{empirical risk}}.$$

To find optimal γ minimizes L_P

- \blacktriangleright Optimize over γ using the above equation directly
- Evaluate the derivative of $L_P(\gamma, s, s')$ by computing covariance

Example: Mean Estimation

 γ^{\ast} depends on other parameters of the problem in a non-trivial manner

$$\gamma^* = \begin{cases} +\infty, & \text{if } \frac{\sigma_Z^2}{n} \in [0, \frac{\sigma_0^2}{2}), \text{ (high-SNR)} \\ \frac{\|\mu - \mu_0\|^2 + d\sigma_0^2/2}{d(2\sigma_Z^2/n - \sigma_0^2)\sigma_0^2}, & \text{if } \frac{\sigma_Z^2}{n} \in [\frac{\sigma_0^2}{2}, \infty). \text{ (low-SNR)} \end{cases}$$

• $\frac{\sigma_Z^2}{n}$ only depends on S, can be interpreted as normalized noise

- $\blacktriangleright~\sigma_0^2$ and $\| \pmb{\mu} \pmb{\mu}_0 \|^2$ captures the confidence and bias of prior knowledge
- high-SNR regime, high-quality training samples, discarding prior distribution and employing standard ERM
- \blacktriangleright low-SNR regime, where we should incorporate knowledge from both training samples and prior, optimal γ depends on everything
- $\blacktriangleright \ \, {\rm If} \ \mu_0=\mu \ {\rm and} \ \sigma_0^2={\tt 0}, \ \gamma^*={\tt 0}$

Example: Linear Regression

- ▶ Training data $S = \{(\mathbf{x}_i, y_i)\}_{i=1}^n$, with $\mathcal{X} = \mathbb{R}^d$ and $\mathcal{Y} = \mathbb{R}$
- \blacktriangleright Data is generated using true weights $W^* \in \mathbb{R}^d$ with additive noise,

$$Y_i = X_i \cdot W^* + arepsilon_i, \quad arepsilon \sim \mathcal{N}(0, \sigma_arepsilon^2).$$

• Mean-squared loss
$$\ell(w, z) = (y - x \cdot w)^2$$

- Gaussian prior $\pi(w) = \mathcal{N}(0, \sigma_0^2 I_d)$
- $(\gamma, \mathcal{N}(0, \sigma_0^2 I_d), L_E(w, s))$ -Gibbs algorithm is Gaussian

$$P_{W|S}^{\gamma}(\boldsymbol{w}|S) \sim \mathcal{N}\Big(\boldsymbol{\Sigma}^{-1} \boldsymbol{X}^{\top} \boldsymbol{Y}, \frac{n}{2\gamma} \boldsymbol{\Sigma}^{-1}\Big),$$

with $\Sigma \triangleq X^{\top}X + \frac{n}{2\sigma_0^2\gamma}I_d$, and $X \in \mathbb{R}^{n \times d}$, $Y \in \mathbb{R}^n$ are the matrix form of the training data.

Simulation of Linear Regression

Low SNR regime, n = 10 and $\sigma_{\varepsilon}^2 = 3$; high SNR regime, n = 100 and $\sigma_{\varepsilon}^2 = 1$.

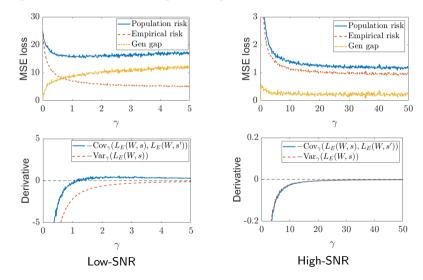


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Asymptotic Behavior of Generalization Error

- Can we show something for ERM by letting $\gamma \to \infty$?
 - Previous upper bound has order $\mathcal{O}(\frac{\gamma}{n})$
- Asymptotic normality of Gibbs algorithm
 - Single-well case: there exists a unique $W^*(S)$

$$W^*(S) = \underset{w \in \mathcal{W}}{\operatorname{arg\,min}} L_E(w, S).$$

► If
$$H^*(S) \triangleq \nabla^2_w L_E(w, S)|_{w=W^*(S)}$$
 is invertible [Hwang, 1980],
 $P^{\gamma}_{W|S} \to \mathcal{N}(W^*(S), \frac{1}{\gamma}H^*(S)^{-1})$

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Asymptotic Behavior of MLE

Maximum likelihood estimates (MLE) in the asymptotic regime $n \to \infty$.

- n i.i.d. training samples generated from distribution P_Z
- ▶ Fit training data with distribution family $f(z_i|m{ heta})$, $m{ heta} \in \mathbb{R}^p$
- $\blacktriangleright \ P_Z = f(\cdot | \boldsymbol{\theta}^*) \text{ for } \boldsymbol{\theta}^* \in \mathcal{W}$
- log-loss $\ell(w, z) = -\log f(z|w)$

As $\gamma \rightarrow \infty,$ Gibbs algorithm converges to ERM algorithm (MLE),

$$\hat{\mathcal{W}}_{\mathrm{ML}} \triangleq rgmax_{oldsymbol{ heta} \in \mathcal{W}} \sum_{i=1}^n \log f(Z_i|oldsymbol{ heta}).$$

Compute $I_{SKL}(W; S)$ using Gaussian approximation

$$\overline{\operatorname{gen}}(P_{W|S}^{\infty},P_S)=\frac{d}{n}.$$

Connection to Model Selection

- K candidate models M_1, M_2, \ldots, M_K
- Each model M_k is characterized by parametric probabilistic model $P_k(\mathbf{z}|\boldsymbol{\theta}_k)$ and prior $\pi_k(\boldsymbol{\theta}_k)$
- ▶ log likelihood as the loss function $\ell_{\log}(w, z) \triangleq -\log P(z|w)$

How to select the optimal model?

- ► Information Criteria for Model Selection
 - ► Akaike Information Criterion (AIC)
 - Bayesian Information Criterion (BIC)

Akaike Information Criterion (AIC)

AIC selects the model that minimizes population risk:

$$\arg\min_{k} \operatorname{KL}(P_{Z} \| P_{k}(\boldsymbol{z} | \hat{\boldsymbol{\theta}}_{\mathrm{ML}}^{(k)})) = \arg\min_{k} \mathbb{E}_{P_{Z}} \big[-\log P_{k}(Z | \hat{\boldsymbol{\theta}}_{\mathrm{ML}}^{(k)}) \big].$$

AIC approximates it using empirical risk and generalization error

$$AIC = \arg\min_{k} L_{E}(\hat{\theta}_{ML}^{(k)}, S) + \overline{gen}(\hat{\theta}_{ML}^{(k)}, P_{Z}).$$

In classic regime where $n
ightarrow \infty$, and certain regularization conditions

AIC =
$$\arg\min_{k} L_{E}(\hat{\theta}_{\mathrm{ML}}^{(k)}, S) + \frac{p}{n}$$

Bayesian Information Criterion (BIC)

BIC selects the model that maximizes marginal likelihood:

$$m_k(\boldsymbol{z}^n) \triangleq \int P_k(\boldsymbol{z}^n|\boldsymbol{\theta}_k) \, \pi_k(\boldsymbol{\theta}_k) \, d\boldsymbol{\theta}_k,$$

which is equivalent to maximizing posterior probability $P(M_k | z^n)$.

BIC =
$$\arg\min_{k} -\frac{1}{n}\log m_{k}(z^{n})$$

= $\arg\min_{k} L_{E}(\hat{\theta}_{ML}^{(k)}, S) + \frac{p_{k}\log n}{2n},$

where Laplace approximation is applied as $n \to \infty$.

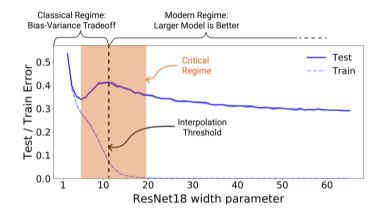
Comparison between AIC and BIC

$$\begin{aligned} \text{AIC} &= \arg\min \ L_E(\hat{\theta}_{\text{ML}},S) + \frac{p}{n} \\ \text{BIC} &= \arg\min \ L_E(\hat{\theta}_{\text{ML}},S) + \frac{p\log n}{2n}. \end{aligned}$$

- AIC minimizes population risk (optimal prediction performance)
- ► BIC maximizes the marginal likelihood (identifying the true model)
- ▶ BIC imposing a larger penalty for more complex models.



Double-descent in Over-parameterized Regime



- When $p \leq n$, the classical \cup -shaped curve is valid.
- When $p \ge n$, test loss can decrease again.

Challenges in Over-parameterized regime

Asymptotic normality (AIC) and Laplace Approximation (BIC) do not hold in this new regime!

There are some efforts to extend these information criteria:

- ▶ Akaike's Information Corrected Criterion (AICC), fixed p, small n
- ▶ Widely applicable BIC (WBIC), singular Hessian matrix

More recent work trying to demystify double-descent

- ▶ Neural Tangent Kernel (NTK), lazy training
- ► Random feature model
- ► Mean-field approach

Marginal likelihood of Gibbs algorithm

Recall the information risk minimization for motivating the Gibbs algorithm.

$$\min_{P_{W|S}} \mathbb{E}_{P_{W|S=s}} \left[L_E(W,s) \right] + \frac{1}{\gamma} \mathrm{KL}(P_{W|S=s} \| \pi) = -\frac{1}{\gamma} \log V(s,\gamma).$$

If we adopt log-loss function $\ell(w, \boldsymbol{z}) = -\log P(\boldsymbol{z}|w)$, and set $\gamma = n$

$$\begin{aligned} -\frac{1}{\gamma}\log V(s,\gamma) &= -\frac{1}{n}\log\int \pi(w)e^{-nL_{\mathcal{E}}(w,s)}dw\\ &= -\frac{1}{n}\log\int \pi(w)P(\boldsymbol{z}^{n}|w)dw\\ &= -\frac{1}{n}\log m(\boldsymbol{z}^{n})\end{aligned}$$

Gibbs based Information Criteria

Gibbs-based AIC:

$$\mathrm{AIC}^+ \triangleq L_{\mathsf{E}}(\hat{\mathcal{W}}_{\mathrm{Gibbs}}, \mathbf{z}^n) + \frac{1}{n} \mathrm{I}_{\mathrm{SKL}}(\mathcal{P}^*_{\hat{\mathcal{W}}|S}, \mathcal{P}_S).$$

Gibbs-based BIC:

$$BIC^{+} \triangleq L_{E}(\hat{W}_{Gibbs}, \boldsymbol{z}^{n}) + \frac{1}{n} KL(P_{W|S=\boldsymbol{z}^{n}}^{*} \| \pi),$$

$$BIC^{-} \triangleq \mathbb{E}_{\pi} [L_{E}(W, \boldsymbol{z}^{n})] - \frac{1}{n} KL(\pi \| P_{W|S=\boldsymbol{z}^{n}}^{*}).$$

We can show that in the classic regime where p is fixed and $n \to \infty$, they all reduce back to their classical forms.

H. Chen, Y. Bu, G. W. Wornell, "Gibbs-Based Information Criteria and the Over-Parameterized Regime," in *Proc. Interna*tional Conference on Artificial Intelligence and Statistics (AISTATS), 2024.

Random Feature Model

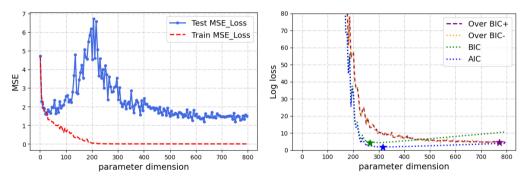
The output of **Random Feature (RF) model** with input data $x \in \mathbb{R}^d$ is

$$g(\mathbf{x}) \triangleq \sum_{j=1}^{p} f\left(\frac{\langle \mathbf{x}, \mathbf{F}_{j} \rangle}{\sqrt{d}}\right) \mathbf{w}_{j} = f\left(\frac{\mathbf{x}^{\top} \mathbf{F}}{\sqrt{d}}\right) \mathbf{w},$$

- ▶ Two-layer neural network with i.i.d Gaussian weights $F \in \mathbb{R}^{d \times p}$ in the first layer, only the second layer is trainable
- ► f() is the non-linear activation function
- \blacktriangleright The dimensionality of input data d is not entangled with number of parameters p

Experiment

Evaluating the BIC⁺ and BIC⁻ using n = 200 samples in RF models



- ▶ We observe **Double-descent** in population risk for RF model
- ► Our Gibbs-based BICs prefer over-parameterized models

- Provide information criteria for the Gibbs algorithm, with different information measures as the penalty terms.
- ► Generalize our information-theoretic analysis to over-parameterized random feature.
- ► The mismatch between marginal likelihood (BIC) and generalization error (AIC) in the over-parameterized setting, which highly depends on the prior distributions.

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Generalization of Transfer Learning

- ▶ Source data set $D^s = \{Z_i^s\}_{i=1}^m$, generated from P_{D^s}
- ▶ Target data set $D^t = \{Z_j^t\}_{j=1}^n$, generated from P_{D^t}
- ▶ The empirical risk of source and target task

$$L_E(w, d^s) \triangleq \frac{1}{m} \sum_{j=1}^m \ell(w, z_j^s), \qquad L_E(w, d^t) \triangleq \frac{1}{n} \sum_{j=1}^n \ell(w, z_j^t).$$

► The population risk of the target task

$$L_P(w, P_{D^t}) \triangleq \mathbb{E}_{P_{D^t}}[L_E(w, D^t)].$$

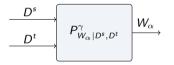
▶ Expected Transfer Generalization Error

$$\overline{\operatorname{gen}}(P_{W|D^s,D^t},P_{D^s},P_{D^t}) \triangleq \mathbb{E}_{P_{W,D^s,D^t}}[L_P(W,P_{D^t}) - L_E(W,D^t)]$$

Transfer Learning: α -Weighted ERM

► Output hypothesis w_{α} is trained by minimizing a convex combination of the source and target task empirical risks [Ben-David et al., 2010], for $\alpha \in [0, 1]$

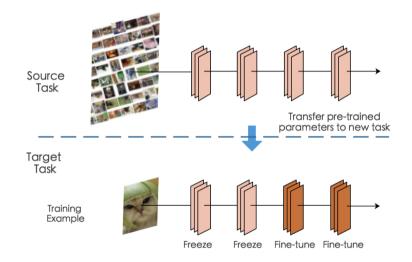
$$L_{E}(w_{\alpha},d^{s},d^{t}) = (1-\alpha)L_{E}(w_{\alpha},d^{s}) + \alpha L_{E}(w_{\alpha},d^{t})$$



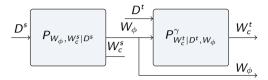
• α -weighted Gibbs algorithm generalizes the α -weighted-ERM by considering the $(\gamma, \pi(w_{\alpha}), L_E(w_{\alpha}, d^s, d^t))$ -Gibbs algorithm

$$\mathcal{P}^{\gamma}_{W_{lpha}|D^{s},D^{t}}(w_{lpha}|d^{s},d^{t})=rac{\pi(w_{lpha})e^{-\gamma L_{E}(w_{lpha},d^{s},d^{t})}}{V_{lpha}(d^{s},d^{t},\gamma)}.$$

Transfer Learning: Two-stage ERM



Two-stage-ERM Transfer Learning



▶ **First Stage:** Learn shared feature extractor $w_\phi \in W_\phi$

$$[W_{\phi}, W_c^s] = \arg\min_w L_E^{S1}(w, d^s).$$

▶ Second Stage: Freeze W_{ϕ} , and learn target-specific hypothesis w_c^t

$$W_c^t = \operatorname*{arg\,min}_{w_c} L_E^{S2}([W_\phi, w_c], d^t)$$

Expected Transfer Generalization Error

Theorem

The expected transfer generalization error of the α -weighted Gibbs algorithm is given by

$$\overline{\operatorname{gen}}_{\alpha}(P_{D^s}, P_{D^t}) = \frac{\operatorname{I}_{\operatorname{SKL}}(W_{\alpha}; D^t | D^s)}{\alpha \gamma}$$

Theorem

The expected transfer generalization error of the two-stage Gibbs algorithm is given by

$$\overline{\operatorname{gen}}_{\beta}(P_{D^s}, P_{D^t}) = \frac{\operatorname{I}_{\operatorname{SKL}}(W_c^t; D^t | W_{\phi})}{\gamma}$$

Y. Bu*, G. Aminian*, L. Toni, M. R. Rodrigues, G. W. Wornell. "Characterizing and Understanding the Generalization Error of Transfer Learning with Gibbs Algorithm," in *Proc. International Conference on Artificial Intelligence and Statistics* (AISTATS) 2022.

Maximum likelihood estimates

- \blacktriangleright *n* i.i.d. target samples, *m* i.i.d. source samples
- ▶ Fit training data with distribution family f(z|w), $w = (w_{\phi}, w_c) \in \mathbb{R}^d$, $w_c \in \mathbb{R}^{d_c}$
- $P_{Z^t} = f(\cdot | \boldsymbol{w}^*)$ for $\boldsymbol{w}^* \in \mathcal{W}$
- log-loss $\ell(w, z) = -\log f(z|w)$

	Standard target ERM	$lpha ext{-weighted ERM}$	Two-stage ERM
$\overline{\mathrm{gen}}$	$\mathcal{O}(\frac{d}{n})$	$\mathcal{O}(\frac{d}{m+n})$	$\mathcal{O}(\frac{d_c}{n})$

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Conclusion

- ► Connect **operational quantity** in learning theory (generalization error, marginal likelihood) with different **information measures** for Gibbs algorithm
- ▶ Demonstrate the versatility of our approach in multiple applications
 - Optimal Inverse temperature
 - ► Gibbs-based BIC for over-parameterized model selection
 - ► Gibbs based-transfer learning
- ► Our Gibbs-based analysis provides an information-theoretic **framework** for understanding generalization behavior in modern machine learning, still a lot to be explored!

Thank you for your attention!

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