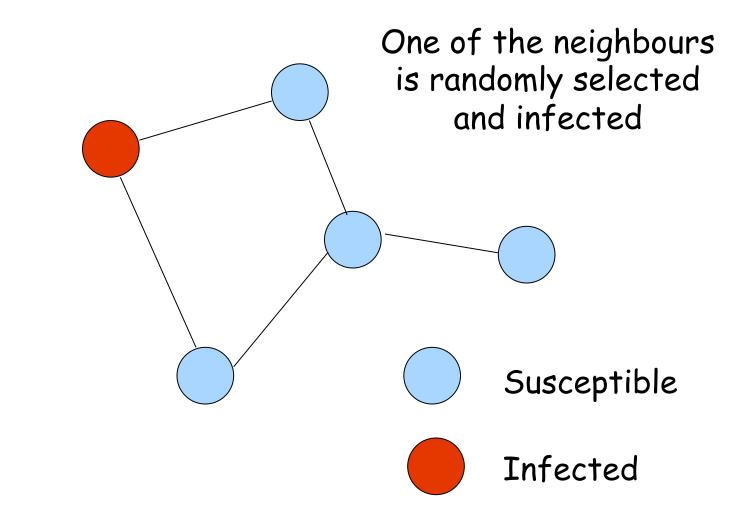
#### Performance Evaluation

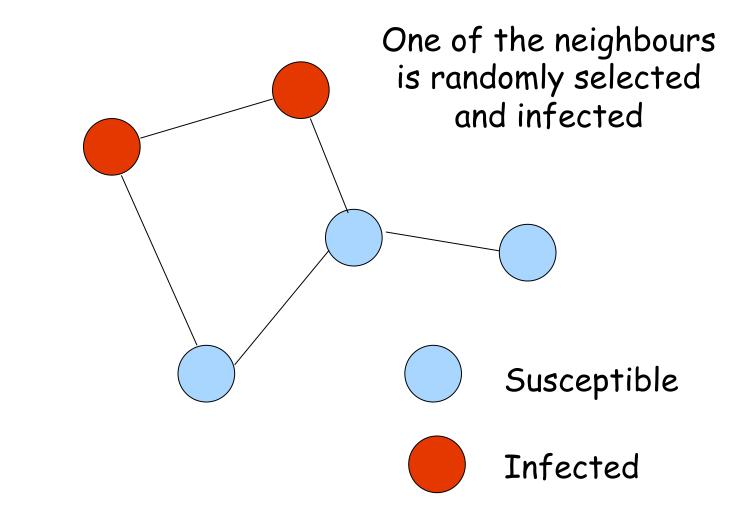
#### **Lecture 2: Epidemics**

Giovanni Neglia INRIA – EPI Maestro 12 December 2012

### Epidemics on a graph: SI model



### Epidemics on a graph: SI model



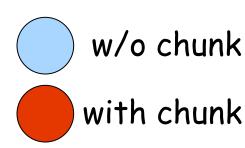
# Any interest for Computer Networks?

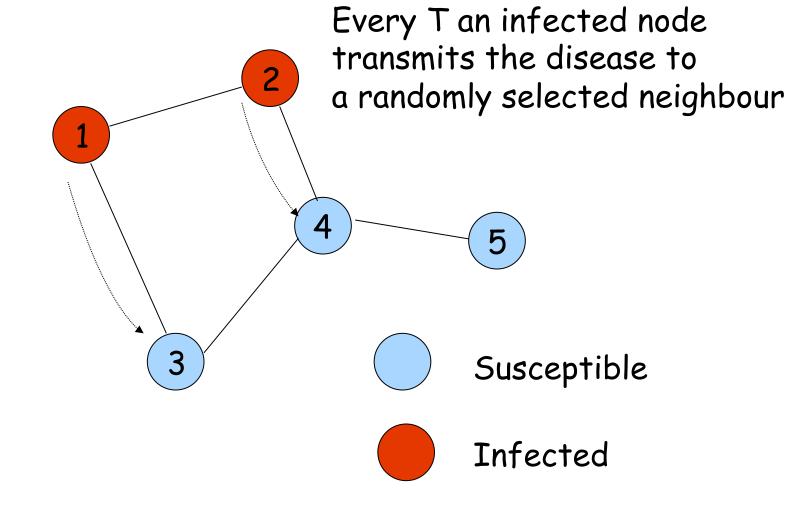
**Flooding** 

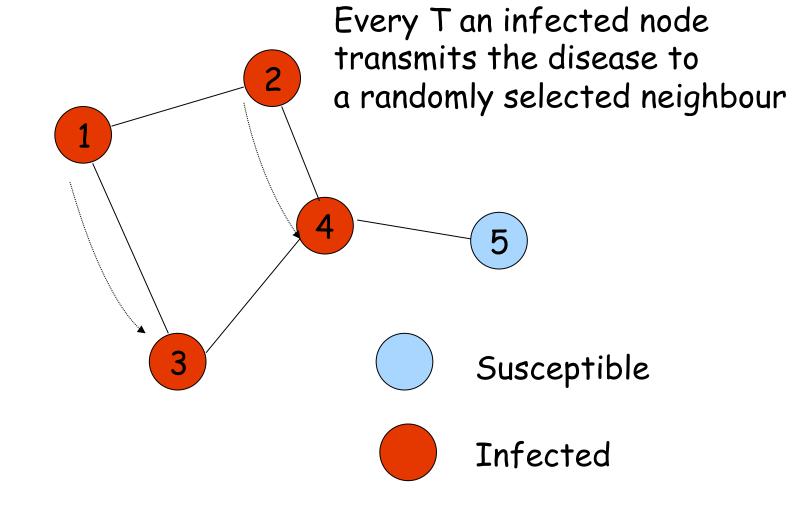
• Epidemic Routing in Delay Tolerant Networks

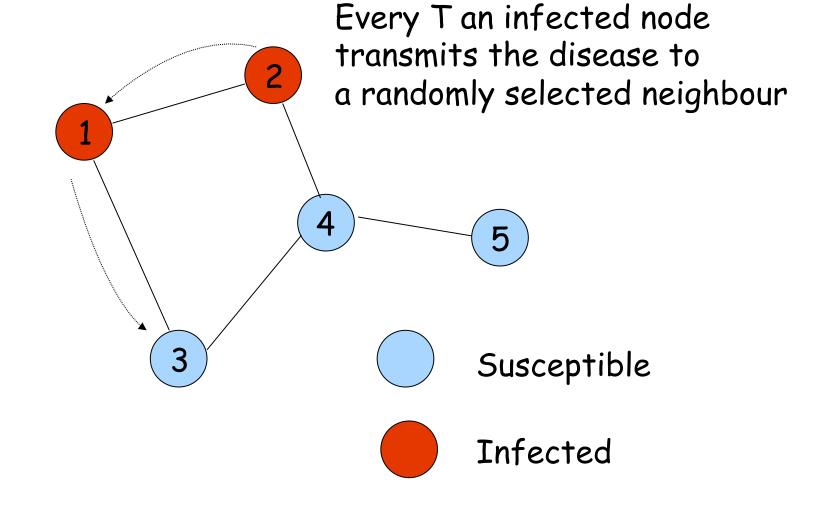
Chunk distribution in a P2P streaming system (push algorithms)

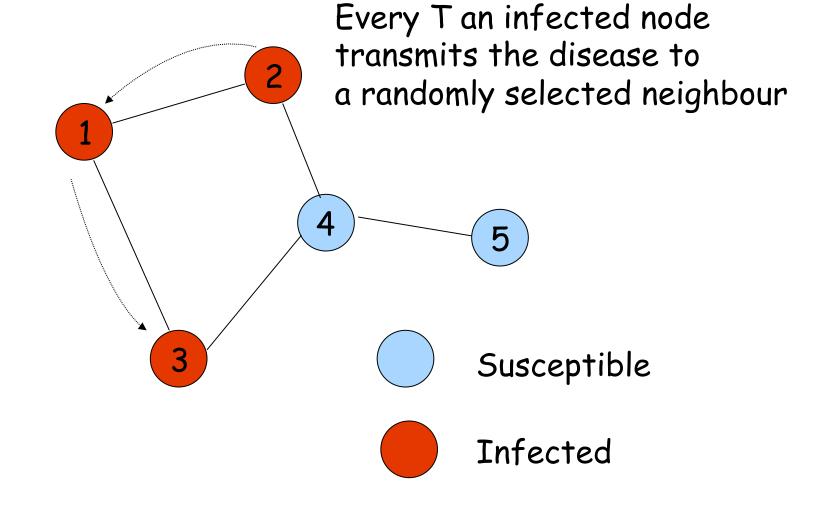
A copy of the chunk is pushed to a randomly selected neighbour







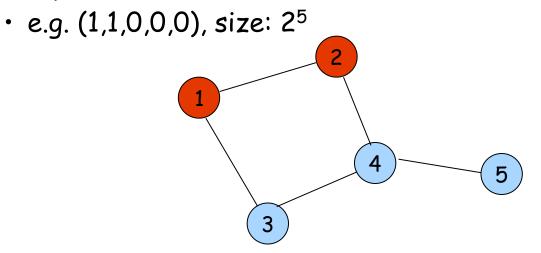




# How do you model it?

#### A Markov Chain

 System state at time k is a vector specifying if every node is infected (1) or not (0)



Probability transitions among states

e.g. Prob((1,1,0,0,0)->(1,1,1,0,0))=1/4

#### Asynchronous behaviour

3

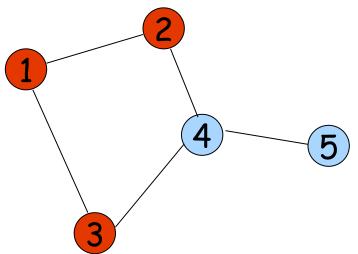
4

Every infected node transmits the disease on each of its links according to a Poisson Process with rate β

5

#### How to model it?

- A Continuous-time Markov process/Chain (C-MC)
  - System state at time t is a vector specifying if every node is infected (1) or not (0)
  - Rate transitions between state pairs
    - e.g, q((1,1,1,0,0)->(1,1,1,1,0))=2β



# What to study and how

- $\square$  P the transition matrix (2<sup>N</sup>x2<sup>N</sup>)
- Transient analysis
  - $\odot \pi(k+1)=\pi(k)P$ ,
  - $\odot \pi(k+1)=\pi(0)P^{k+1}$ ,
- Stationary distribution (equilibrium)

  - If the Markov chain is irreducible and aperiodic
  - Computational cost:
    - $O((2^N)^3)$  if we solve the system
    - O(K M) where M is the number of non-null entries in P if we adopt the iterative procedure (K is the number of iterations), in our case  $M=O((2^N)^2)$

#### Similar for C-MC

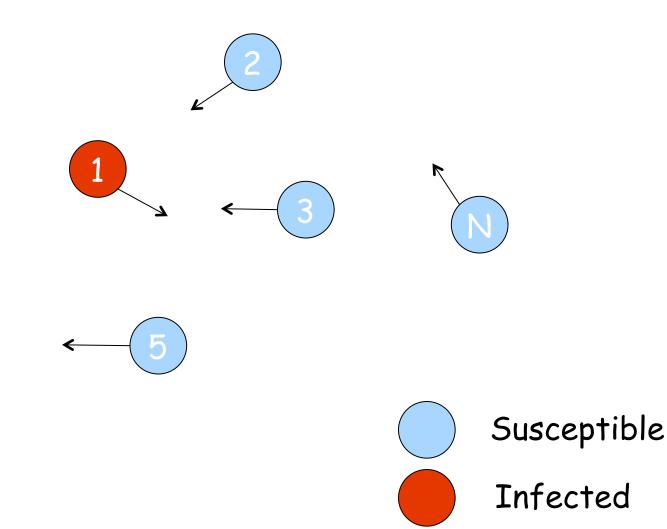
- □ Stationary distribution (equilibrium)  $\circ \pi = \pi P$ , D-MC
  - πQ=0, C-MC

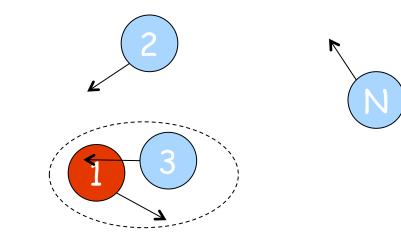
Transient analysis

 π(k+1)=π(k)P, D-MC
 dπ(t)/dt=π(t)Q, C-MC

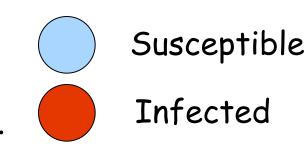
#### Outline

- Limit of Markovian models
- Mean Field (or Fluid) models
  - exact results
  - extensions to graphs
  - applications

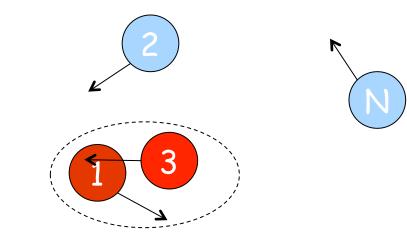




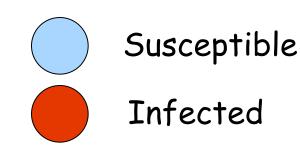
At each slot there is a probability p that two given nodes meet. Assume meetings to be independent.



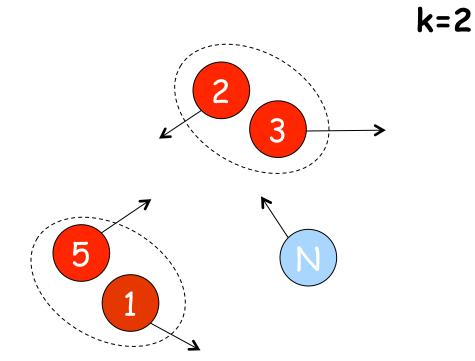
k=1



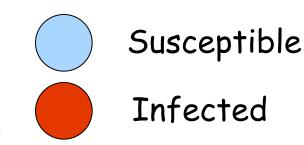
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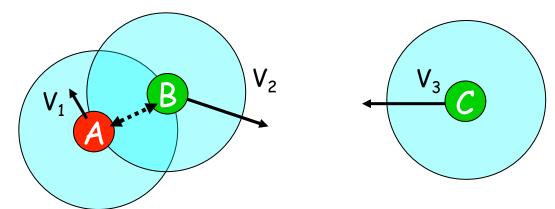
k=1



At each slot there is a probability p that two given nodes meet. Assume meetings to be independent.



#### Delay Tolerant Networks (a.k.a. Intermittently Connected Networks)



mobile wireless networks
no path at a given time instant between two nodes
because of power contraint, fast mobility dynamics
maintain capacity, when number of nodes (N) diverges
Fixed wireless networks: C = O(sqrt(1/N)) [Gupta99]
Mobile wireless networks: C = O(1), [Grossglauser01]
a really challenging network scenario
No traditional protocol works

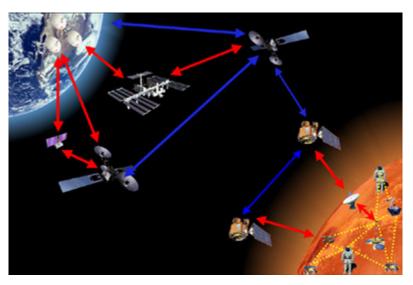
#### Some examples





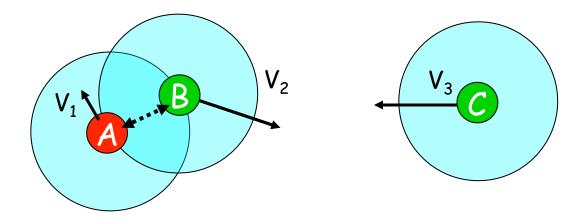


- Network for disaster relief team
- Military battle-field network



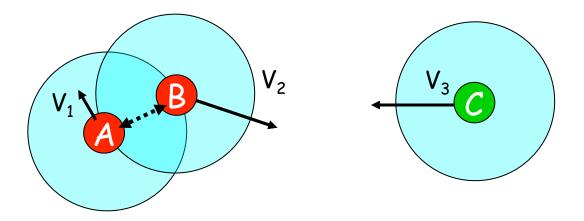
#### Inter-planetary backbone

# **Epidemic Routing**



Message as a disease, carried around and transmitted

# **Epidemic Routing**



Message as a disease, carried around and transmitted

Store, Carry and Forward paradigm

# How do you model it?

#### A Markov Chain

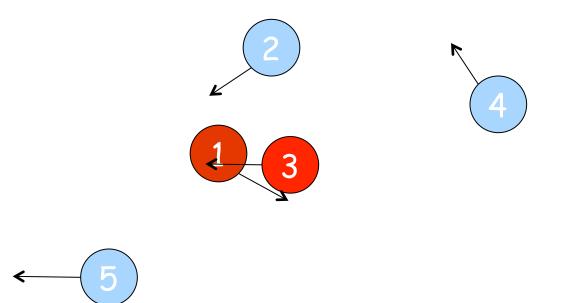
 System state at time k is a vector specifying if every node is infected (1) or not (0)

• e.g. (1,0,1,0,0), size: 2<sup>5</sup> 3 4 5

Probability transitions among states

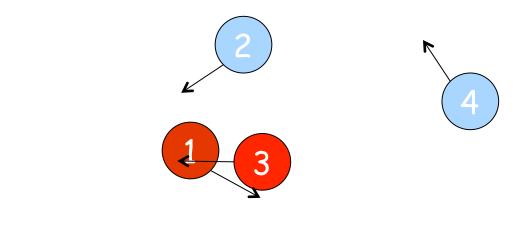
• e.g. Prob((1,0,1,0,0)->(1,1,1,0,0))=?

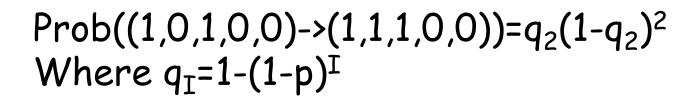
# **Transition probabilities** Prob((1,0,1,0,0)->(1,1,1,0,0))=?



At slot k, when there are I(=I(k)) infected nodes, the prob. that node 2 gets infected is:  $q_I=1-(1-p)^I$ 

# **Transition probabilities** Prob((1,0,1,0,0)->(1,1,1,0,0))=?





# What to study and how

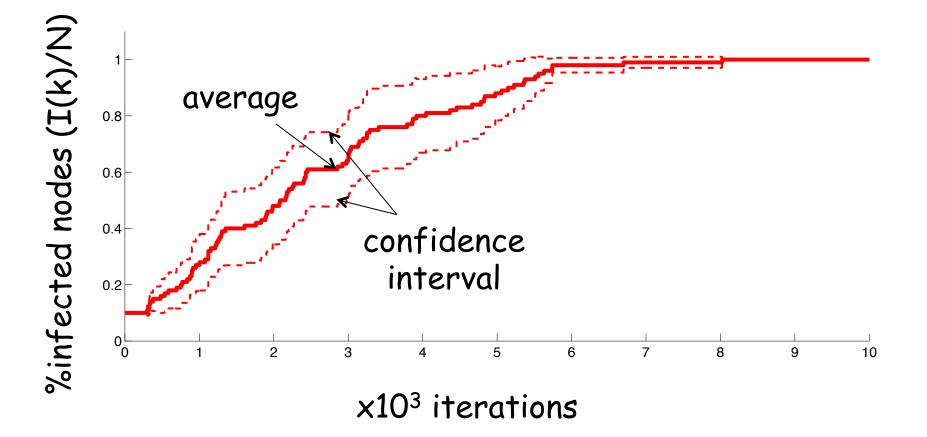
- $\square$  P the transition matrix (2<sup>N</sup>x2<sup>N</sup>)
- Transient analysis
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  - If the Markov chain is irreducible and aperiodic
  - Computational cost:
    - $O((2^N)^3)$  if we solve the system
    - O(K M) where M is the number of non-null entries in P if we adopt the iterative procedure (K is the number of iterations), in our case  $M=O((2^N)^2)$

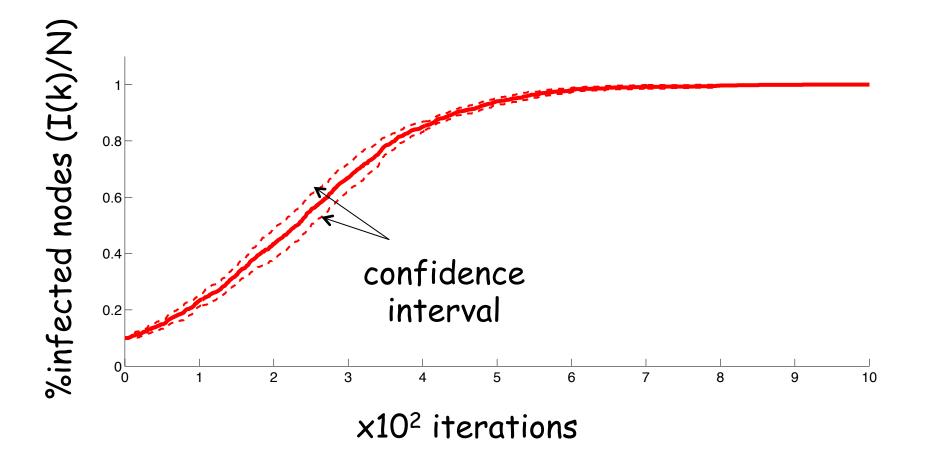
# Can we simplify the problem?

- all the nodes in the same state (infected or susceptibles) are equivalent
- If we are interested only in the number of nodes in a given status, we can have a more succinct model
  - o state of the system at slot k: I(k)
  - it is still a MC
  - $O \operatorname{Prob}(I(k+1)=I+n | I(k)=I) = C_{N-I}^{n} q_{I}^{n} (1-q_{I})^{N-n-I}$ 
    - $(I(k+1)-I(k) | I(k)=I) \sim Bin(N-I,q_I)$
    - q<sub>I</sub>=1-(1-p)<sup>I</sup>

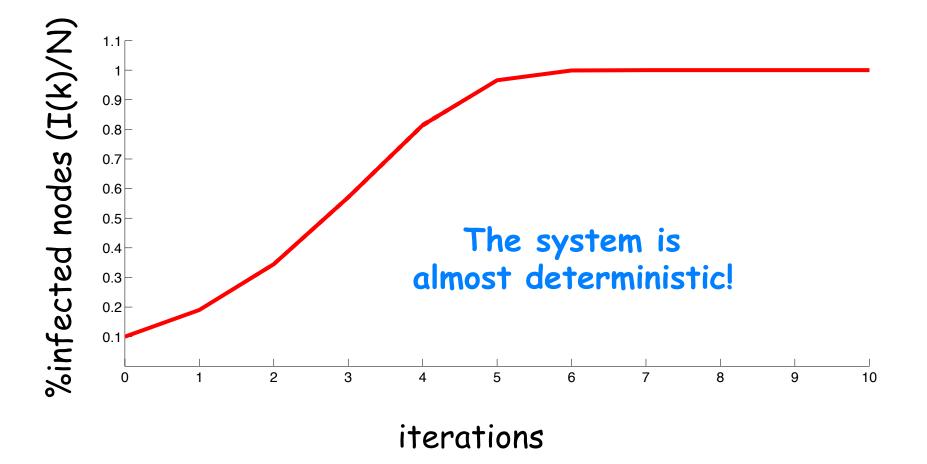
# Some numerical examples p=10<sup>-4</sup>, N=10, I(0)=N/10, 10 runs



# Some numerical examples p=10<sup>-4</sup>, N=100, I(0)=N/10, 10 runs



#### Some numerical examples p=10<sup>-4</sup>, N=10000, I(0)=N/10, 10 runs



#### Summary

- For a large system of interacting equivalent objects, the Markov model can be untractable...
- but a deterministic description of the system seems feasible in terms of the empirical measure (% of objects in each status)
  - intuition: kind of law of large numbers
- Mean field models describe the deterministic limit of Markov models when the number of objects diverges



- i<sup>(N)</sup>(k), fraction of infected nodes at time k
   Solve
  - di(t)/dt=i(t)(1-i(t)), with i=i<sub>0</sub>
    - Solution: i(t)=1/((1/i<sub>0</sub>-1) e<sup>-+</sup>+1)
- If  $i^{(N)}(0)=i_0$ ,  $i^{(N)}(k) \approx i(k p_0/N)=1/((1/i_0-1) \exp(-k p_0/N)+1)$  $=1/((1/i_0-1) \exp(-k N p)+1)$

## Outline

Limit of Markovian models
Mean Field (or Fluid) models

- exact results
- extensions to graphs
- applications

#### References

- Results here for discrete time Markov Chains
  - Benaïm, Le Boudec "A Class of Mean Field Interaction Models for Computer and Communication Systems", LCA-Report-2008-010
- A survey with pointers to continuous time Markov processes and links to stochastic approximation and propagation of chaos
  - Ch. 2 of Nicolas Gast's PhD thesis "Optimization and Control of Large Systems, Fighting the Curse of Dimensionality"

## Necessary hypothesis: Objects' Equivalence

**Π** π(k+1)=π(k)P

□ A state  $\sigma = (v_1, v_2, ..., v_N), v_j \in V(|V|=V, finite)$ 

E.g. in our example V={0,1}

 $\Box$  P is invariant under any label permutation  $\varphi$ :

• 
$$P_{\sigma,\sigma'}=Prob((v_1,v_2,...v_N)-(u_1,u_2,...u_N))=$$
  
 $Prob((v_{\phi(1)},v_{\phi(2)},...v_{\phi(N)})-(u_{\phi(1)},u_{\phi(2)},...u_{\phi(N)}))$ 

#### Some notation and definitions

- $\Box X_n^{(N)}(k)$ : state of node n at slot k
- $\square$   $M_{(N)}(k)$ : occupancy measure of state v at slot k
  - $M_{u}^{(N)}(k) = \sum_{n} \mathbf{1}(X_{n}^{(N)}(k) = v)/N$
  - SI model:  $M_2^{(N)}(k) = I^{(N)}(k) / N = i^{(N)}(k)$ ,

 $M_1^{(N)}(k) = S^{(N)}(k) / N = s^{(N)}(k) = 1 - i^{(N)}(k)$ 

- $\square M^{(N)}(k) = (M_1^{(N)}(k), M_2^{(N)}(k), \dots M_V^{(N)}(k))$

 $\Box f^{(N)}(m) = E[M^{(N)}(k+1) - M^{(N)}(k)]M^{(N)}(k) = m]$ 

Drift or intensity, it is the mean field

- - SI model:  $(1-i^{(N)}(k),i^{(N)}(k))$

# Other hypotheses

 $\Box$  Intensity vanishes at a rate  $\varepsilon(N)$ 

 $- \operatorname{Lim}_{N \rightarrow \infty} \mathbf{f}^{(N)}(\mathbf{m}) / \epsilon(N) = \mathbf{f}(\mathbf{m})$ 

- Second moment of number of object transitions per slot is bounded
  - #transitions< $W^{N}(k)$ ,
    - $\mathsf{E}[W^{\mathsf{N}}(\mathsf{k})^2 | \mathbf{M}^{(\mathsf{N})}(\mathsf{k}) = \mathbf{m}] < c \mathsf{N}^2 \varepsilon(\mathsf{N})^2$
- Drift is a smooth function of m and 1/N
  - $f^{(N)}(m)/\epsilon(N)$  has continuous derivatives in m and in 1/N on  $[0,1]^{V} \times [0,\beta]$ , with  $\beta > 0$

#### Convergence Result

 $\Box$  Define <u>**M**</u><sup>(N)</sup>(t) with t real, such that

- $\underline{\mathbf{M}}^{(N)}(k \epsilon(N)) = \mathbf{M}^{(N)}(k)$  for k integer
- $\underline{\mathbf{M}}^{(N)}(t)$  is affine on [k  $\varepsilon(N),(k+1)\varepsilon(N)$ ]
- Consider the Differential Equation
  - $-d\mu(t)/dt=f(\mu)$ , with  $\mu(0)=m_0$

Theorem

– For all T>O, if  $\mathbf{M}^{(N)}(0) \rightarrow \mathbf{m}_0$  in probability (/mean square) as  $N \rightarrow \infty$ , then  $\sup_{0 \le t \le T} ||\underline{\mathbf{M}}^{(N)}(t) - \mu(t)|| \rightarrow 0$  in probability (/ mean square)

#### Convergence of random variables

The sequence of random variables X<sup>(N)</sup> converges to X in probability if

- for all  $\delta$ >0 Lim<sub>N→∞</sub> Prob(|X<sup>(N)</sup> - X|> $\delta$ )=0

The sequence of random variables X<sup>N</sup> converges to X in mean square if

$$- \text{Lim}_{N \to \infty} \mathbb{E}[|X^{(N)} - X|^2] = 0$$

Convergence in mean square implies convergence in probability

- Assumptions' check
  - Nodes are equivalent
  - Intensity vanishes at a rate  $\varepsilon(N)$  $f^{(N)}(m) = E[M^{(N)}(k+1) - M^{(N)}(k)]M^{(N)}(k) = m]$  $M_{2}^{(N)}(k)=I^{(N)}(k)/N=i^{(N)}(k), M_{1}^{(N)}(k)=1-M_{2}^{(N)}(k)$  $(I^{(N)}(k+1)-I^{(N)}(k) | I^{(N)}(k)=I) \sim Bin(N-I_{q_T}) =>$  $E[I^{(N)}(k+1)-I^{(N)}(k) | I^{(N)}(k)=I] = q_T(N-I)$  $E[i^{(N)}(k+1)-i^{(N)}(k)|i^{(N)}(k)=i] = (1-i) q_T$ =  $(1-i)(1-(1-p)^{iN}) \rightarrow (1-i)$  when N diverges!

Out of the impasse: introduce a scaling for p

- If  $p^{(N)}=p_0/N^a a>1 => (1-i)(1-(1-p^{(N)})^{i})>0$
- Consider a=2
  - $(1-i)(1-(1-p^{(N)})^{iN}) \sim (1-i) i p_0/N$  (for N large)
- ε(N)=p<sub>0</sub>/N
- $f_2(m) = f_2((s,i)) = s i = i (1-i)$
- Lesson to keep: often we need to introduce some parameter scaling

- Assumptions' check
  - Nodes are equivalent
  - Intensity vanishes at a rate  $\epsilon(N)=p_0/N$
  - Second moment of number of object transitions per slot is bounded

#transitions<W<sup>N</sup>(k),

 $\mathsf{E}[\mathsf{W}^{\mathsf{N}}(\mathsf{k})^2 | \mathbf{M}^{(\mathsf{N})}(\mathsf{k}) = \mathbf{m}] < c \mathsf{N}^2 \varepsilon(\mathsf{N})^2$ 

 $W^{N}(k)$ =#trans. ~ Bin(N-I(k),q\_I)

 $E[W^{N}(k)^{2}]=((N-I(k))q_{I})^{2} + (N-I(k))q_{I}(1-q_{I})$ is in  $O(N^{2} \epsilon(N)^{2})$ 

- Assumptions' check
  - Nodes are equivalent
  - ✓ Intensity vanishes at a rate  $\epsilon(N)=p_0/N$
  - Second moment of number of object transitions per slot is bounded
  - $\checkmark$  Drift is a smooth function of **m** and 1/N
    - $f_2^{(N)}(\mathbf{m})/\epsilon(N) =$ =(1-i) (1 - (1-(p\_0/N^2))^{i} N)/(p\_0/N)
    - continuous derivatives in i and in 1/N (not evident)

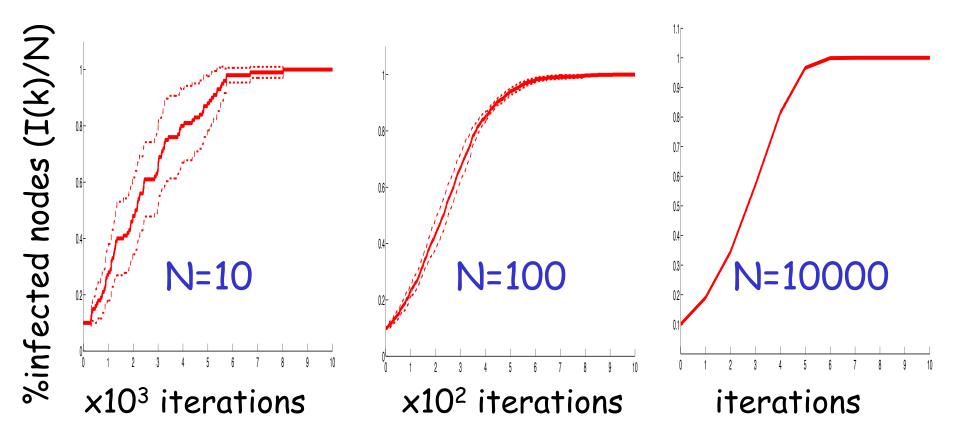
# Practical use of the convergence result

#### Theorem

- For all T>O, if  $\mathbf{M}^{(N)}(0) \rightarrow \mathbf{m}_0$  in probability (/mean square) as  $N \rightarrow \infty$ , then  $\sup_{0 \le t \le T} ||\underline{\mathbf{M}}^{(N)}(t) - \mu(t)|| \rightarrow 0$  in probability (/ mean square)
- Where  $\mu(t)$  is the solution of  $d\mu(t)/dt=f(\mu)$ , with  $\mu(0)=m_0$
- $\square \mathbf{M}^{(N)}(0) = \mathbf{m}_{0}, \ \mathbf{M}^{(N)}(k) = \underline{\mathbf{M}}^{(N)}(k\varepsilon(N)) \approx \boldsymbol{\mu}(k\varepsilon(N))$

- □  $f_2(\mathbf{m})=f_2((s,i))=i(1-i)$ □  $d\mu_2(t)/dt=f_2(\mu_2(t))=\mu_2(t)(1-\mu_2(t)),$ with  $\mu_2(0)=\mu_{0,2}$ 
  - Solution:  $\mu_2(t)=1/((1/\mu_{0,2}-1)e^{-t}+1)$
- □ If  $i^{(N)}(0)=i_0$ ,  $i^{(N)}(k) \approx \mu_2 (k\epsilon(N))=1/((1/i_0-1) \exp(-k p_0/N)+1)$  $=1/((1/i_0-1) \exp(-k N p)+1)$

# Back to the numerical examples p=10<sup>-4</sup>, I(0)=N/10, 10 runs



# Advantage of Mean Field

- □ If  $i^{(N)}(0)=i_0$ ,  $i^{(N)}(k) \approx \mu_2 (k\epsilon(N))=1/((1/i_0-1) \exp(-k p_0/N)+1)$ =1/((1/i\_0-1) exp(-k N p)+1)
  - solved for each N with negligible computational cost
- In general: solve numerically the solution of a system of ordinary differential equations (size = #of possible status)
  - simpler than solving the Markov chain