



2. Mathematical Preliminaries

Mathematical Preliminaries



- **Random Variables**
- **Linear Transformations**
- **Eigenvalues and Eigenvectors**
- **Orthonormal Transformations**
- **Matrix Inversion**

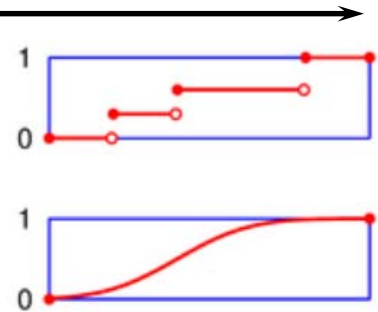
Random Variables

In statistical pattern recognition,
a pattern (input to a PR system) is
a d-dimensional *feature
vector* (*random variable*)

$$\mathbf{X} = (X_1, X_2, \dots, X_d)^T$$

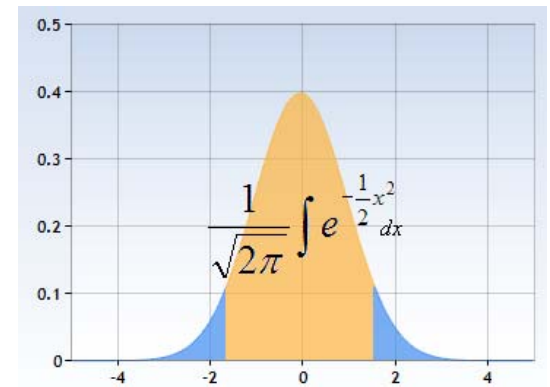
Random Variables

Is fully characterized by its (cumulative) distribution function:



$$F_{\mathbf{X}}(x_1, \dots, x_n) = Pr\{X_1 \leq x_1, \dots, X_d \leq x_d\}$$

or density function:



$$p_{\mathbf{X}}(x_1, \dots, x_d) = \partial^d F(x_1, \dots, x_d) / \partial x_1 \dots \partial x_d$$

Random Variables

- In pattern recognition, we deal with random vectors drawn from different classes
- Conditional density of class i (L classes):

$$p(\mathbf{x}|\omega_i) = p_i(\mathbf{x}), \quad (i = 1, \dots, L)$$

- Unconditional density function of \mathbf{x} (mixture density function):

$$p(\mathbf{x}) = \sum_{i=1}^L P_i p_i(\mathbf{x})$$

P_i is a priori probability of class i

Random Variables

- A posteriori probability of w_i given \mathbf{x} (Bayes theorem):

$$P(\omega_i|\mathbf{x}) = q_i(\mathbf{x}) = \frac{P_i p_i(\mathbf{x})}{p(\mathbf{x})}$$

Random Variables

- Expected vector (mean):

$$\mathbf{X} = (X_1, X_2, \dots, X_d)^T$$

$$M = E(\mathbf{X}) = (\mu_1, \mu_2, \dots, \mu_d)^T$$

Random Variables

- **Covariance matrix (indicates the dispersion of the distribution):**

$$\Sigma = E\{(\mathbf{X} - M)(\mathbf{X} - M)^T\} = \begin{bmatrix} c_{11} & \dots & c_{1d} \\ & \vdots & \\ c_{d1} & \dots & c_{dd} \end{bmatrix}$$

$$c_{ii} = \sigma_i^2, \quad c_{ij} = \rho_{ij}\sigma_i\sigma_j$$

σ_i^2 is the variance, σ_i is the standard deviation of \mathbf{X}_i

ρ_{ij} is the correlation coefficient between \mathbf{X}_i and \mathbf{X}_j

Random Variables

$$\begin{aligned}\Sigma &= E\{\mathbf{X}\mathbf{X}^T\} - E\{\mathbf{X}\}M^T - ME\{\mathbf{X}^T\} + MM^T \\ &= E\{\mathbf{X}\mathbf{X}^T\} - MM^T = S - MM^T\end{aligned}$$

$E\{\mathbf{X}\mathbf{X}^T\} = S$ is the autocorrelation function

Random Variables

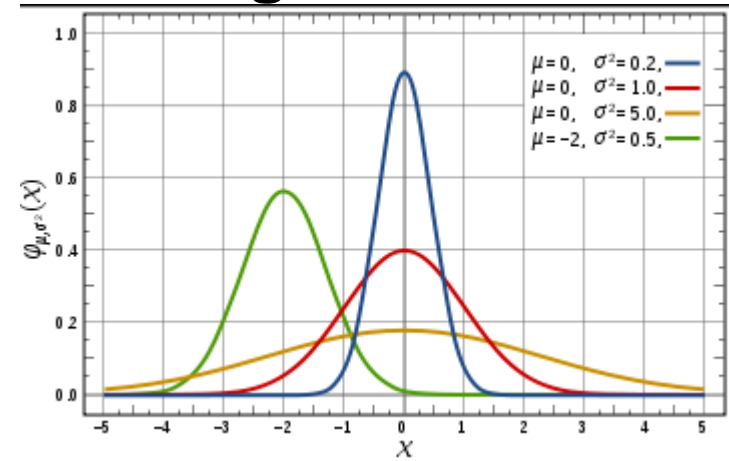
Gaussian (normal) distribution: describes data that cluster around a mean or average

1-d

$$p(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{(x - \mu)^2}{2\sigma^2}\right\}$$

d>=1

$$p(\mathbf{x}|M, \Sigma) = \frac{1}{(\sqrt{2\pi})^d \sqrt{|\Sigma|}} \exp\left\{-\frac{1}{2} \underbrace{(\mathbf{x}-M)^T \Sigma^{-1} (\mathbf{x}-M)}_{\text{distance function}}\right\}$$



Random Variables

Normal distribution:

- is uniquely characterized by the expected vector and covariance matrix
- The assumption of normality is a reasonable approximation for many real data sets
- However, normality should not be assumed without good justification
- [More about normal distribution](#)

Linear Transformations

$$\mathbf{y} = \mathbf{A}\mathbf{x} \Rightarrow \begin{aligned} y_1 &= a_{11}x_1 + \dots + a_{1d}x_d \\ &\vdots \\ y_d &= a_{d1}x_1 + \dots + a_{dd}x_d \end{aligned}$$

$$\Leftrightarrow \begin{bmatrix} y_1 \\ \vdots \\ y_d \end{bmatrix} = \begin{bmatrix} a_{11} & \dots & a_{1d} \\ & \vdots & \\ a_{d1} & \dots & a_{dd} \end{bmatrix} \begin{bmatrix} x_1 \\ \vdots \\ x_d \end{bmatrix}$$

Linear Transformations

$$M_{\mathbf{y}} = E\{\mathbf{y}\} = E\{A\mathbf{x}\} = AE\{\mathbf{x}\} = AM_{\mathbf{x}}$$

$$\Rightarrow M_{\mathbf{y}} = AM_{\mathbf{x}}$$

$$\Sigma_{\mathbf{y}} = E\left\{ \underbrace{(\mathbf{y} - M_{\mathbf{y}})}_{A(\mathbf{x} - M_{\mathbf{x}})} \underbrace{(\mathbf{y} - M_{\mathbf{y}})^T}_{(\mathbf{x} - M_{\mathbf{x}})^T A^T} \right\} = A\Sigma_{\mathbf{x}}A^T$$

Eigenvalues and Eigenvectors

$$Q\phi_i = \lambda_i\phi_i$$

\swarrow
 $d \times d$

$\phi_i, i = 1, \dots, d$
- eigenvectors
 λ_i - eigenvalues

Write for all i :

$$Q[\phi_1, \dots, \phi_d] = [\lambda_1\phi_1, \dots, \lambda_d\phi_d]$$
$$= [\phi_1, \dots, \phi_d] \begin{bmatrix} \lambda_1 & & 0 \\ & \dots & \\ 0 & & \lambda_d \end{bmatrix}$$

Eigenvalues and Eigenvectors

$$Q[\phi_1, \dots, \phi_d] = [\lambda_1 \phi_1, \dots, \lambda_d \phi_d]$$

$$= [\phi_1, \dots, \phi_d] \begin{bmatrix} \lambda_1 & & 0 \\ & \dots & \\ 0 & & \lambda_d \end{bmatrix}$$

eigenvector matrix

$$Q\Phi = \Phi\Lambda$$

eigenvalue matrix

Eigenvalues and Eigenvectors

$$Q\Phi = \Phi\Lambda$$

The eigenvectors corresponding to two different eigenvalues are orthogonal:

$$\phi_i^T \phi_j = 0, \lambda_i \neq \lambda_j$$

$$\Rightarrow \Phi^T \Phi = I$$

$$\Rightarrow \Phi^T = \Phi^{-1}$$

$$\Phi^T Q \Phi = \Lambda$$

Eigenvalues and Eigenvectors

$$\Phi^T \Sigma_X \Phi = \Lambda$$

$$\mathbf{y} = \Phi^T \mathbf{x}$$

$$\Phi^T \Sigma_X \Phi = \Sigma_Y = \begin{bmatrix} \lambda_1 & & 0 \\ & \dots & \\ 0 & & \lambda_d \end{bmatrix}$$

$\rho_{ij} = 0$

Eigenvalues and Eigenvectors

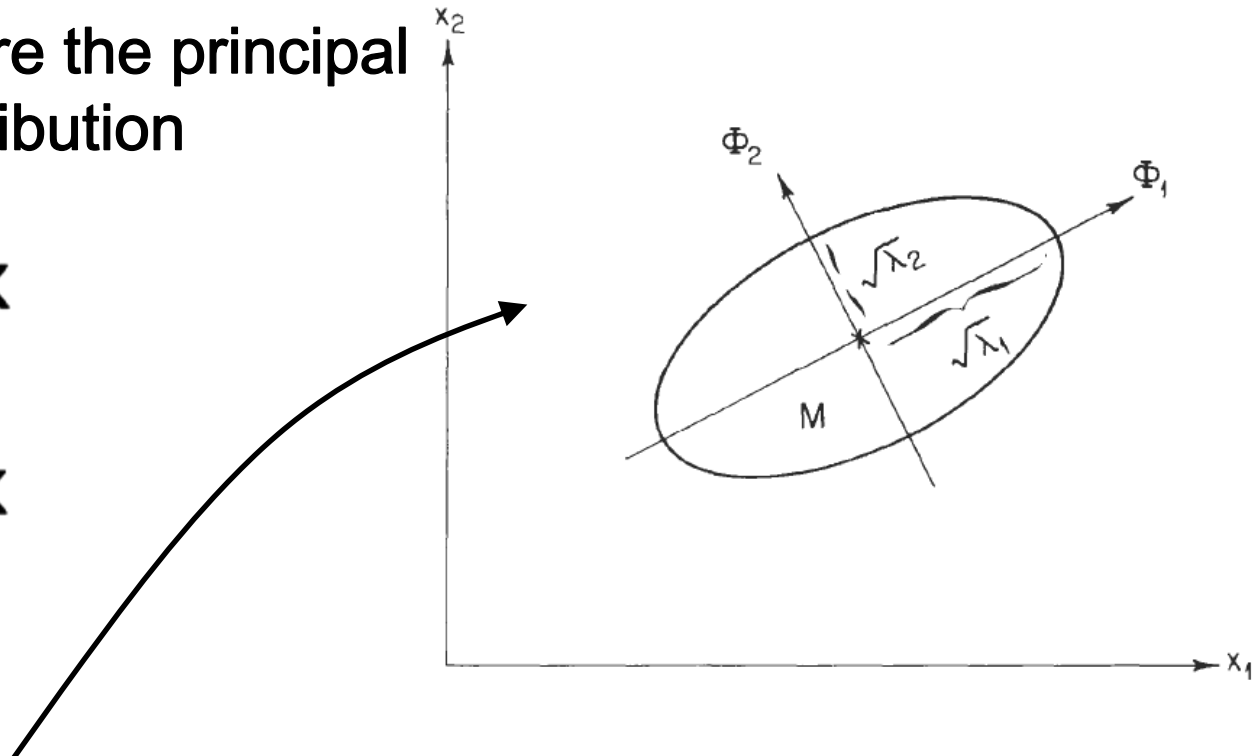
→ There is no correlation in the transformed space:

$$\mathbf{y} = \Phi^T \mathbf{x} = \begin{bmatrix} \phi_1^T \\ \vdots \\ \phi_d^T \end{bmatrix} \mathbf{x}$$

Eigenvalues and Eigenvectors

Eigenvectors are the principal axis of the distribution

$$\begin{aligned}
 y_1 &= \phi_1^T \mathbf{x} \\
 &\vdots \\
 y_d &= \phi_d^T \mathbf{x}
 \end{aligned}$$



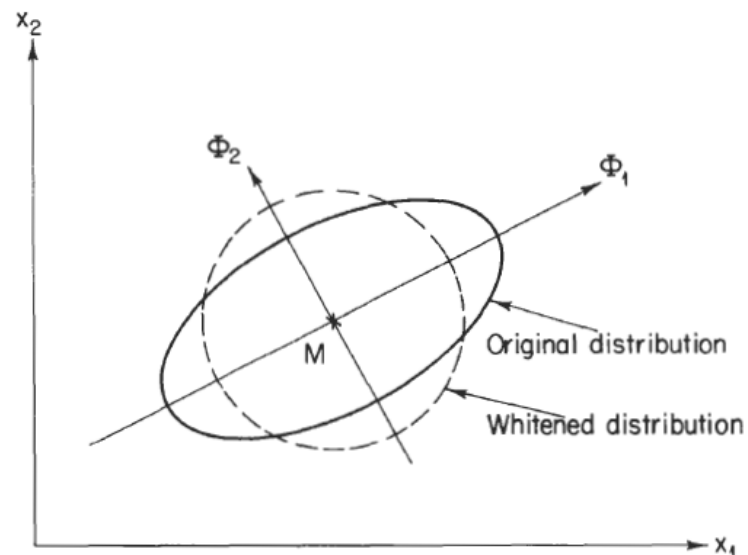
- We rotate the coordinate system
- There is no correlation between \$y_1\$ and \$y_2\$
- \$\lambda_1\$ gives the variance of \$y_1\$, \$\lambda_2\$ gives the variance of \$y_2\$

Whitening

After applying the orthonormal transformation, we can add another transformation $\Lambda^{-1/2}$ that will make the covariance matrix equal to I :

$$\mathbf{y} = \Lambda^{-1/2} \Phi^T \mathbf{x}$$

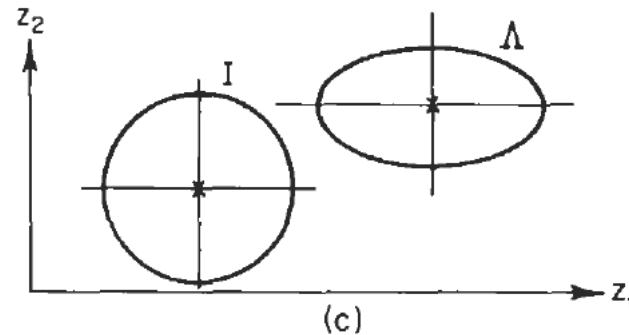
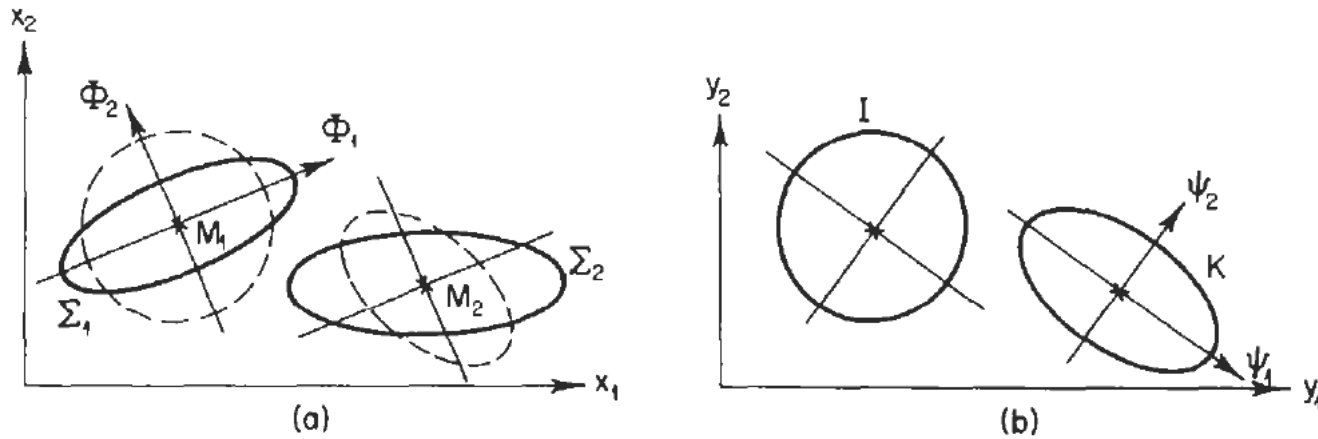
$$\begin{aligned} \Sigma_{\mathbf{y}} &= \Lambda^{-1/2} \Phi^T \Sigma_{\mathbf{x}} \Phi \Lambda^{-1/2} \\ &= \Lambda^{-1/2} \Lambda \Lambda^{-1/2} = I \end{aligned}$$



Purpose: to change the scales of principal components in proportion to $1/\sqrt{\lambda_i}$

Simultaneous Diagonalization

Goal: diagonalize Σ_1 and Σ_2 simultaneously by a linear transformation:



Orthonormal Transformations

$$\mathbf{y} = A\mathbf{x} \quad AA^T = I$$

Properties:

a) Eigenvalues are invariant

$$\Sigma_{\mathbf{x}}\Phi = \Phi\Lambda$$

$$\Sigma_{\mathbf{y}}\Psi = \Psi K$$

$$\mathbf{y} = A\mathbf{x} \Rightarrow A\Sigma_{\mathbf{x}}A^T = \Sigma_{\mathbf{y}}$$

$$\Rightarrow A\Sigma_{\mathbf{x}}A^T\Psi = \Psi K$$

Orthonormal Transformations

$$\Rightarrow \Sigma_{\mathbf{X}} \underbrace{A^T \Psi}_{\Phi} = \underbrace{A^T \Psi}_{\Phi} \underbrace{K}_{\Lambda}$$

$$\Sigma_{\mathbf{X}} \Phi = \Phi \Lambda$$

Orthnormal Transformations

b) Euclidean distance is invariant

$$\mathbf{x}^T \mathbf{x} = \sum_{i=1}^d x_i^2$$

$$\mathbf{y}^T \mathbf{y} = (\mathbf{A}\mathbf{x})^T (\mathbf{A}\mathbf{x}) = \mathbf{x}^T \mathbf{A}^T \mathbf{A}\mathbf{x} = \mathbf{x}^T \mathbf{x}$$

→ Distance in Y-space is the same as in X-space

Trace of Covariance Matrix

$\text{tr } \Sigma_{\mathbf{X}} =$ Sum of diagonal terms of $\Sigma_{\mathbf{X}}$

Let's look at $\Phi \Lambda \Phi^T$

$$\begin{aligned} \text{tr } \Phi \Lambda \Phi^T &= \text{tr } \Lambda \Phi^T \Phi = \text{tr } \Lambda \\ &= \sum_{i=1}^d \lambda_i = \text{tr } \Sigma_{\mathbf{X}} \end{aligned}$$

→ The trace of $\Sigma_{\mathbf{X}}$ is the summation of all eigenvalues and is invariant under any orthonormal transformation

Determinant of Covariance Matrix

$$|AB| = |A||B|$$

$$\begin{aligned} |\Sigma_{\mathbf{x}}| &= |\Phi \Lambda \Phi^T| = |\Phi| |\Lambda| |\Phi^T| = \\ |\Lambda| |\Phi^T| |\Phi| &= |\Lambda| |\Phi^T \Phi| = |\Lambda| |I| = |\Lambda| = \prod_{i=1}^d \lambda_i \end{aligned}$$

→ The determinant of $\Sigma_{\mathbf{x}}$ is equal to the product of all eigenvalues and is invariant under any orthonormal transformation

Rank of Covariance Matrix

Rank of $\Sigma_{\mathbf{X}}$ is equal to the number of nonzero eigenvalues

$$\Sigma_{\mathbf{X}} = \Phi \Lambda \Phi^T = \sum_{i=1}^d \lambda_i \phi_i \phi_i^T$$

Relation between $|S|$ and $|\Sigma|$:

$$|S| = |\Sigma| (1 = M^T \Sigma^{-1} M)$$

Matrix Inversion



Introduction to matrix inversion: 1 and 2

Generalized Inverse (Pseudo Inverse)

$$QQ^{-1} = I$$

$$Q = \sum_{i=1}^d \lambda_i \phi_i \phi_i^T$$

$$Q^{-1} = \sum_{j=1}^d \frac{1}{\lambda_j} \phi_j \phi_j^T$$

$$I = QQ^{-1} = \sum_{i=1}^d \sum_{j=1}^d \frac{1}{\lambda_j} \lambda_i \phi_i \phi_i^T \phi_j \phi_j^T = \sum_{i=1}^d \frac{\lambda_i}{\lambda_i} \phi_i \phi_i^T = \sum_{i=1}^d \phi_i \phi_i^T$$

Generalized Inverse (Pseudo Inverse)

However, if matrix is singular \rightarrow some eigenvectors are zero

$$Q = \sum_{i=1}^r \lambda_i \phi_i \phi_i^T \quad r < d, \lambda_i = 0 \text{ for } i = r + 1, \dots, d$$

Generalized (pseudo) inverse:

$$Q^* = \sum_{j=1}^r \frac{1}{\lambda_j} \phi_j \phi_j^T$$

$$QQ^* = \sum_{i=1}^r \phi_i \phi_i^T \neq I$$

because we skip
 $i = r + 1, \dots, d$

But:

$$QQ^*Q = Q$$