

Microscopic behavior of TCP

Vincent Dumas

Fabrice Guillemin

Philippe Robert

Bert Zwart

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References

- [1] Vincent Dumas, Fabrice Guillemin, and Philippe Robert, A Markovian analysis of additive-increase multiplicative-decrease (AIMD) algorithms, *Advances in Applied Probability* **34** (2002), no. 1, –.
- [2] Fabrice Guillemin, Philippe Robert, and Bert Zwart, AIMD algorithms and exponential functionals, Submitted to the *Annals of Applied Probability* (2002).

1. Introduction

Setting, Assumptions and Objectives

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Setting

- Long TCP connection (Elephant).
- Network is a black box rejecting some packets.
- Analysis of the successive congestion window sizes (W_n) when loss rate α is small.

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- Model for packet losses.

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Objectives

- Prove rigorous convergence results.
- Give explicit expressions of the distributions involved.

Congestion avoidance:

The additive increase and multiplicative decrease scheme

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The additive increase and multiplicative decrease scheme

W_n is the size of the n th congestion window.

If $W_n = x \in \mathbb{N}$,

$$W_{n+1} = \begin{cases} \lfloor \delta x \rfloor & \text{if one of the } x \text{ packets is lost,} \\ x + 1 & \text{otherwise} \end{cases}$$

Stochastic models for packet losses: I

An elementary model: the non correlated case

- Each packet has a probability α of being lost.
- Packet losses are independent.

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The independence assumption

Reasonable: the network evolves rapidly.

Not realistic: losses occur by groups.

2. Non correlated case: Convergence results

(W_n^α) : sequence of successive window sizes.

Transitions of the Markov chain.

If $W_0^\alpha = x$

- $W_1^\alpha = \lfloor \delta x \rfloor$, with probability $1 - \exp(-\alpha x)$.
- $W_1^\alpha = x + 1$, with probability $\exp(-\alpha x)$.

The invariant probability is $\pi^\alpha = (\pi_n^\alpha)$, W_∞^α is some random variable with distribution π^α .

The congestion window size is $\sim 1/\sqrt{\alpha}$.

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Theorem. *If $\lim_{\alpha \rightarrow 0} \sqrt{\alpha} W_0^\alpha = x$, and*

$$W^\alpha(t) = \sqrt{\alpha} W_{\lfloor t/\sqrt{\alpha} \rfloor}^\alpha,$$

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Theorem. If $\lim_{\alpha \rightarrow 0} \sqrt{\alpha} W_0^\alpha = x$, and

$$W^\alpha(t) = \sqrt{\alpha} W_{\lfloor t/\sqrt{\alpha} \rfloor}^\alpha,$$

then $(W^\alpha(t))$ converges in distribution to the Markov process $(\bar{W}(t))$ whose infinitesimal generator is given by

$$\Omega(f)(x) = f'(x) + x (f(\delta x) - f(x))$$

and $\bar{W}(0) = x$.

The embedded Markov chain (V_n^α)

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Theorem. If $\lim_{\alpha \rightarrow 0} \sqrt{\alpha} V_0^\alpha = \bar{v}$, then $(\sqrt{\alpha} V_n^\alpha)$ converges in distribution to the Markov chain (\bar{V}_n) , with $\bar{V}_0 = \bar{v}$ and the transitions

$$\bar{V}_{n+1} \stackrel{\text{dist.}}{=} \delta(\bar{V}_n + \bar{G}_{\bar{V}_n}),$$

where for $x \geq 0$,

$$\mathbb{P}(\bar{G}_x \geq y) = e^{-(xy + y^2/2)}.$$

Convergence diagrams

$$\begin{array}{ccc} \left(\sqrt{\alpha} W_{\lfloor t/\sqrt{\alpha} \rfloor}^{\alpha} \right) & \xrightarrow{t \rightarrow +\infty} & \sqrt{\alpha} W_{\infty}^{\alpha} \\ \alpha \rightarrow 0 \downarrow & & \\ (\overline{W}(t)) & & \end{array}$$

$$\begin{array}{ccc} (\sqrt{\alpha} V_n^{\alpha}) & \xrightarrow{n \rightarrow +\infty} & \sqrt{\alpha} V_{\infty}^{\alpha} \\ \alpha \rightarrow 0 \downarrow & & \\ (\overline{V}_n) & & \end{array}$$

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Theorem. *The above diagrams commute.*

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Theorem. *The above diagrams commute.*

Proof: Exponential estimates of hitting times.

3. Non correlated case: The equilibrium

Proposition. *The following equality in distribution holds*

$$(x + \overline{G}_x)^2 \stackrel{\text{dist.}}{=} 2E_1 + x^2,$$

for $x \geq 0$, where E_1 is an exponentially distributed random variable with parameter 1.

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Consequence for the embedded chain:

$$\overline{V}_{n+1} = \delta (\overline{V}_n + \overline{G}_{\overline{V}_n})$$

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Consequence for the embedded chain:

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$$\begin{aligned} \overline{V}_{n+1}^2 &= \delta^2 (\overline{V}_n + \overline{G}_{\overline{V}_n})^2 \\ &\stackrel{\text{dist.}}{=} \delta^2 (\overline{V}_n^2 + 2E_n) \end{aligned}$$

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Proposition. *The sequence (\overline{V}_n^2) is AR (Auto-Regressive):*

$$\overline{V}_{n+1}^2 = \delta^2 (\overline{V}_n^2 + 2E_n)$$

(E_n) i.i.d. exponential variables with parameter 1.

Invariant measure of the AR process

Theorem. *The distribution of the random variable \bar{V}_∞ is given by*

$$\bar{V}_\infty^2/2 \stackrel{\text{dist.}}{=} \sum_{n=1}^{+\infty} \delta^{2n} E_n \stackrel{\text{dist.}}{=} \int_0^{+\infty} \delta^{2N(s)} ds.$$

where (E_i) are i.i.d. exp. random variables with parameter 1;
 N Poisson process with parameter 1.

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Its density h_δ given by

$$h_\delta(x) = \frac{1}{\prod_{n=1}^{+\infty} (1 - \delta^n)} \sum_{n=1}^{+\infty} \frac{1}{\prod_{k=1}^{n-1} (1 - \delta^{-2k})} \delta^{-2n} x e^{-\delta^{-2n} x^2/2},$$

for $x \geq 0$.

Corollary. The *throughput* of the TCP model is given by

$$\lim_{\alpha \rightarrow 0} \sqrt{\alpha} \mathbb{E}(W_{\infty}^{\alpha}) \stackrel{\text{def.}}{=} \bar{\rho}(\delta) = \frac{1}{\mathbb{E}(\bar{V}_{\infty})} = \sqrt{\frac{2}{\pi}} \prod_{n=1}^{+\infty} \frac{(1 - \delta^{2n})}{(1 - \delta^{2n-1})}.$$

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$$\bar{\rho}(1/2) \sim 1.3098,$$

Corollary. The *throughput* of the TCP model is given by

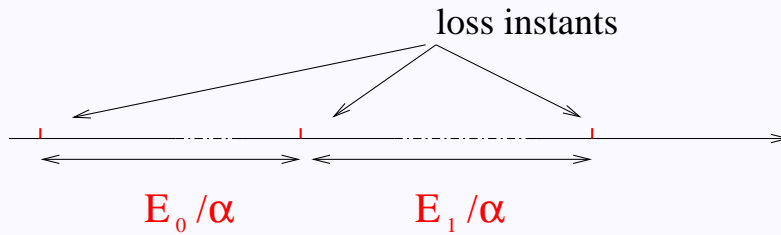
$$\lim_{\alpha \rightarrow 0} \sqrt{\alpha} \mathbb{E}(W_{\infty}^{\alpha}) \stackrel{\text{def.}}{=} \bar{\rho}(\delta) = \frac{1}{\mathbb{E}(\bar{V}_{\infty})} = \sqrt{\frac{2}{\pi}} \prod_{n=1}^{+\infty} \frac{(1 - \delta^{2n})}{(1 - \delta^{2n-1})}.$$

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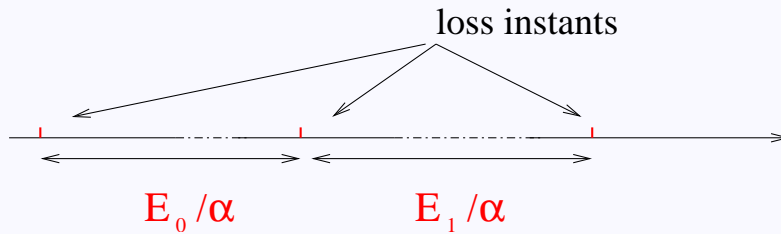
Tail distribution: $\mathbb{P}(\bar{W}_{\infty} \geq x) \sim C \exp(-x^2/\delta^2)$.

4. General loss processes

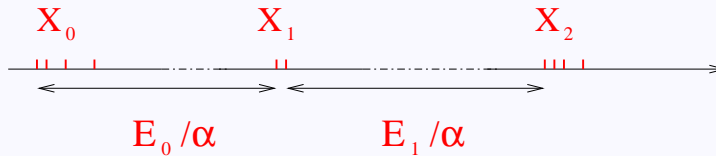
Non correlated model:



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"Real" loss process (Paxson 1997):



Dynamic of congestion window size

$$\bar{V}_{n+1}^2 = \delta^2 (\bar{V}_n^2 + 2E_n)$$

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(X_n) i.i.d. integer valued random variables.

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Equilibrium

$$\bar{V}_\infty^2/2 \stackrel{\text{dist.}}{=} \beta^{X_0} (\bar{V}_\infty^2/2 + E_1)$$

$$\bar{V}_\infty^2/2 \stackrel{\text{dist.}}{=} I = \int_0^{+\infty} \beta^{X(s)} ds,$$

with $\beta = \delta^2$,

$$X(t) = \sum_{k=0}^{N(t)} X_i,$$

N Poisson process with parameter 1.

I is an integral exponential functional of a Lévy process

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with $Y(t) = (X_1 + \cdots + X_{N(t)}) \log \beta$, N is a Poisson process with parameter 1.

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with $Y(t) = (X_1 + \dots + X_{N(t)}) \log \beta$, N is a Poisson process with parameter 1.

When $Y(t) = B(t) - \mu t$, ($B(t)$) Brownian motion and $\mu > 0$, I plays a role in mathematical finance models and statistical physics.

Work by Bertoin, Carmona, Monthus, Petit, Yor (and many others).

The distribution of I

The Laplace transform of I is given by,

$$\mathbb{E}(e^{-\lambda I}) = \sum_{m=0}^{\infty} \frac{(-\lambda)^m}{\prod_{k=1}^m (1 - \mathbb{E}(\beta^k X_1))},$$

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If $X_1 \in \{1, \dots, n+1\}$,

$$1 - \mathbb{E}(u^{X_1}) = (1 - u)(1 - a_1 u) \dots (1 - a_n u),$$

$$\mathbb{E}(e^{-\lambda I}) = \sum_{m=0}^{\infty} \frac{(-\lambda)^m}{(a_1 \beta; \beta)_m \dots (a_n \beta; \beta)_m (\beta; \beta)_m}$$

with the q -factorial $(a; q)_m = (1 - a)(1 - aq) \dots (1 - aq^{m-1})$
 $1 \leq m \leq +\infty$.

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The Laplace transform of I is related to a q -hypergeometric function.

Applications

The density of I is given by

$$\sum_{k=0}^{+\infty} r_k \beta^{-k} e^{-\beta^{-k} x},$$

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1. $\mathbb{P}(X_1 = 1) = p = 1 - \mathbb{P}(X_1 = 2)$.

$$r_k = \frac{1}{(-\beta(1-p); \beta)_{\infty} (\beta; \beta)_{\infty} (1/\beta; 1/\beta)_k} \sum_{m=0}^k \frac{(1/\beta; 1/\beta)_k}{(1/\beta; 1/\beta)_m (1/\beta; 1/\beta)_{k-m}} (-(1-p))^m.$$

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2. $\mathbb{P}(X_1 = n) = p^{n-1}(1-p), n \geq 1.$

$$r_k = \frac{1}{(\beta; \beta)_{\infty}} \frac{(\beta^{k+1} p; \beta)_{\infty}}{(\beta; \beta)_{k-1}}$$

The fractional moments of I

Proposition. For any $s \in \mathbb{R}$, $-s \notin \mathbb{N} - \{0\}$, if $\mathbb{E}(\beta^{(s+1)X_1}) < +\infty$,

$$\mathbb{E}(I^s) = \Gamma(s+1) \prod_{k=1}^{+\infty} \frac{1 - \mathbb{E}(\beta^{(s+k)X_1})}{1 - \mathbb{E}(\beta^{kX_1})},$$

for $u \geq \min(s+1, 0)$.

Proposition. *The asymptotic throughput of an AIMD algorithm with multiplicative decrease factor δ in a correlated loss model is given by*

$$\bar{\rho}_X = \lim_{\alpha \rightarrow 0} \sqrt{\alpha \mathbb{E}(X_1) \mathbb{E}(\bar{W}_\infty^\alpha)} = \sqrt{\frac{2\mathbb{E}(X_1)}{\pi} \prod_{n=1}^{+\infty} \frac{1 - \mathbb{E}(\delta^{2n} X_1)}{1 - \mathbb{E}(\delta^{(2n-1)} X_1)}}.$$

Question: What are the properties of the mapping

$$X \rightarrow \bar{\rho}_X(\delta)?$$

Impact of correlation of the loss process

Definition. [Concave ordering of random variables]
The inequality $X \leq_{cv} Y$ holds when

$$\mathbb{E}(f(X)) \leq \mathbb{E}(f(Y))$$

is true for any non-decreasing concave function on \mathbb{R} .

Impact of correlation of the loss process

Proposition. *The asymptotic throughput $X \rightarrow \bar{\rho}_X / \sqrt{\mathbb{E}(X)}$ is a non-increasing function for the concave order,*

$$X \leq_{cv} Y \quad \text{implies} \quad \frac{\bar{\rho}_X}{\sqrt{\mathbb{E}(X)}} \geq \frac{\bar{\rho}_Y}{\sqrt{\mathbb{E}(Y)}}. \quad (1)$$

In particular, when $\mathbb{E}(X) = \mathbb{E}(Y)$, $X \leq_{cv} Y$ implies $\bar{\rho}_X \geq \bar{\rho}_Y$.

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In particular, when $\mathbb{E}(X) = \mathbb{E}(Y)$, $X \leq_{cv} Y$ implies $\bar{\rho}_X \geq \bar{\rho}_Y$.

Corollary. For a fixed loss rate

- The more variable the loss process, the better the throughput.
- The non correlated model is worst than *any* loss model.

Discussion

- Analogous results for **Finite Maximum Window size**.
- Slow start period. Duration $-\log_2(\alpha)$: negligible at the time scale $1/\sqrt{\alpha}$.
- Variable RTT's. Under weak assumptions, formulas are still valid even with dependence.
- Occurrences of timeouts: $\mathbb{P}(X_1 = +\infty) > 0$.