

On the compromise between burstiness and frequency of events

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Consider a process of events with two types: **good** and **lost**.
 n consecutive events are called a block.

Time may be continuous

but the model will be in discrete-time and ignore actual time intervals between events.

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Metric of interest: given h and n ,

$$P(\text{ the block is "lost" }) = P(> h \text{ losses among } n \text{ events})$$

Usual objective: find the smallest h (redundancy) such that

$$P(> h \text{ losses among } n + h \text{ events}) < \varepsilon .$$

Today's objective: compare two situations

- same event loss probability
- different "burstiness" patterns

Motivation #1: Forward Error Correction

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Forward error correction at the packet level: able to repair up to h lost packets, using h packets of redundancy.



k=8 information packets + **h=4 redundancy packets**



OK



LOST

Motivation #1 (ctd)

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Different **queue management schemes** at routers produce different loss patterns.

Assuming the loss rate is the same: is it better

- to have losses regularly spaced,
- or have losses clustered?

Additional motivation

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Reliability/real time systems:

- n tasks to be executed within a time frame
each one may fail
execute $m = n + k$ of them
“ k -out-of- m ”

Bandwidth reduction in a slotted network:

- frames of n slots \rightarrow frames of h slots
no buffer
probability of overflow?

Pedestrian crossing...

Well known facts...

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Several facts are well known:

- **Variability worsen things** (Folk result)
 - ⇒ the situation with the “most regular arrivals” should be better
- **The independence assumption is optimistic**
 - If the loss events are independent, the block loss probabilities are (much) smaller than if they are correlated
 - ⇒ the situation with the “most independent arrivals” should be better

Investigate the issue with a focus on the **bursts of losses**.

Well known facts...

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Mathematical experiment # 1: The Gilbert model

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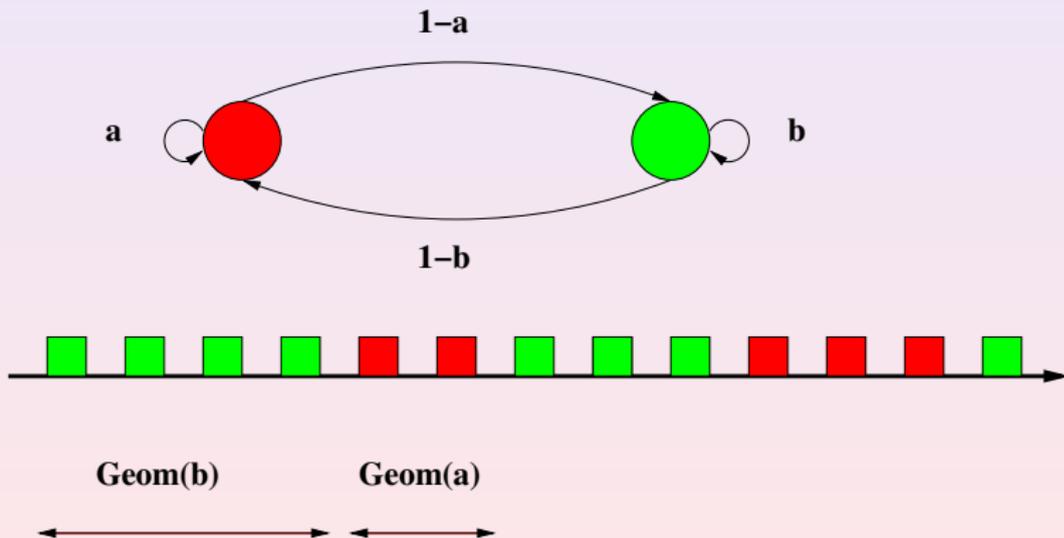
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Assumption: losses occur according to the state of a (two-state) Markov chain.



The Gilbert model (2)

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Gilbert as a Markov-Additive process:

$$L_{m+1} = L_m + \mathbf{1}_{\{X_m = \bullet\}}.$$

$$\begin{aligned} E(z^{L_n}) &= \pi_0 M(z)^n \mathbf{1} \\ &= (\pi_{\bullet}, \pi_{\circ}) \times \begin{pmatrix} az & (1-a)z \\ 1-b & b \end{pmatrix}^n \times \begin{pmatrix} 1 \\ 1 \end{pmatrix}, \end{aligned}$$

where

$$\pi_{\bullet} = \frac{1-b}{2-a-b} \quad \pi_{\circ} = \frac{1-a}{2-a-b}.$$

Gilbert model (3)

Loss Run Length (LRL):

$$\text{LRL} = \frac{1}{1-a}$$

Good Run Length (GRL):

$$\text{GRL} = \frac{1}{1-b}$$

Stationary loss probability:

$$p = \pi_{\bullet} = \frac{\text{LRL}}{\text{LRL} + \text{GRL}}.$$

Problem: with a fixed LRL (or a), the range of p is

$$\left[0, \frac{\text{LRL}}{\text{LRL} + 1}\right).$$

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Skewed Gilbert Model

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Solution: make Good Runs Geometrically distributed on $\{0, 1, \dots\}$ instead of $\{1, 2, \dots\}$.

\implies another Gilbert process with matrix:

$$\begin{pmatrix} 1 - b(1 - a) & b(1 - a) \\ 1 - b & b \end{pmatrix},$$

and

$$\text{LRL} = \frac{1}{1 - a} \quad \text{GRL} = \frac{b}{1 - b} \quad p = \frac{1 - b}{1 - ab}.$$

Now the range of p is $[0, 1]$!

Skewed Gilbert Model

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Comparison Bernoulli/Gilbert

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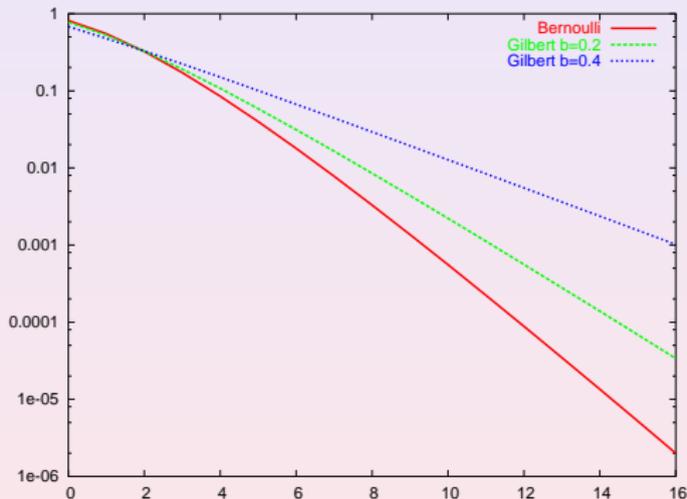
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Loss probability of a block of size $n = h + 16$, depending on h .

Comparison experiments (1)

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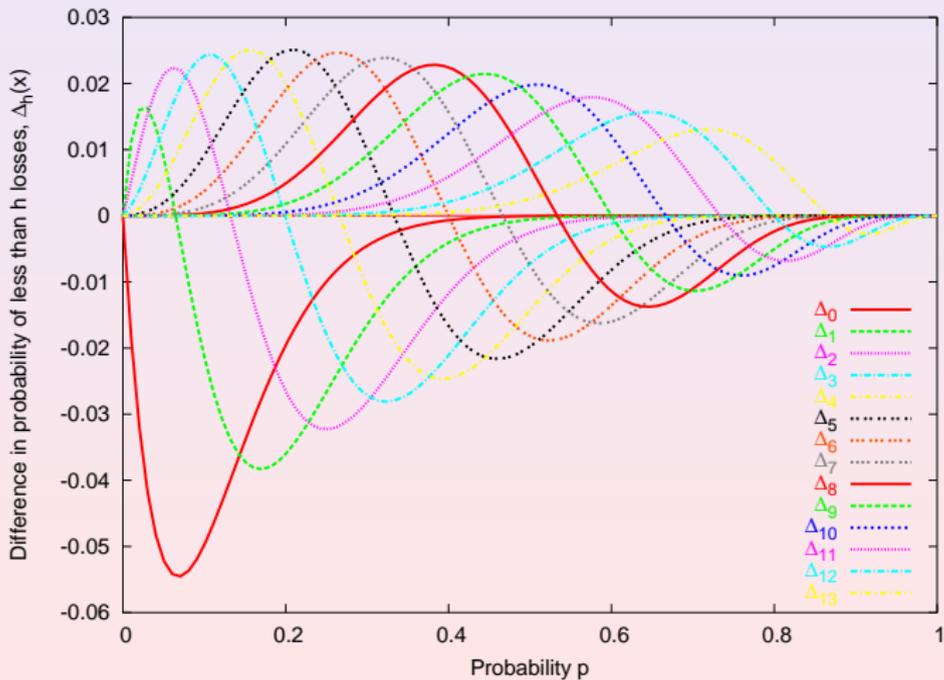
Experiment: consider two cases 1 and 2.

- Fix the Loss Run lengths: $LRL_1 < LRL_2$ ($a_1 < a_2$),
- fix a block length k and a “redundancy” quantity h
- vary the Loss Probability p
- plot the difference:

$$\Delta_h(p) = P(\text{block saved in case 1}) \\ - P(\text{block saved in case 2})$$

Comparison experiments (2)

h grows from 0 (left, red) to 13 (right, yellow).



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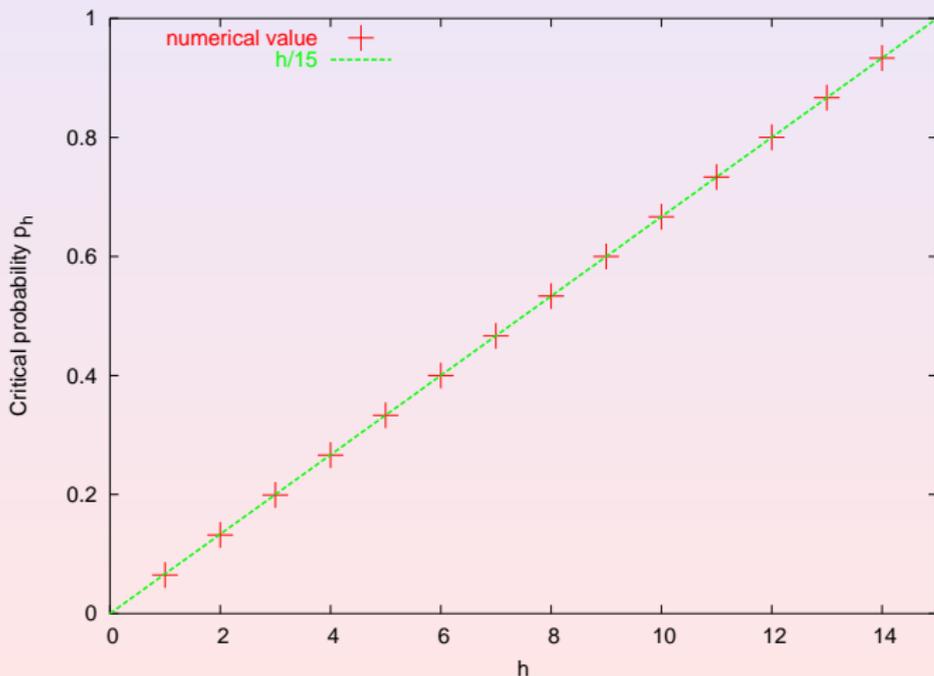
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Comparison experiments (3)

Let p_h be the value at which $\Delta_h(p_h) = 0$.



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Work in progress

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Empirical finding: when n is large,

$$x_h \sim \frac{h}{n-1} .$$

How to prove it?

If the loss rate is $p = h/n$,

$$P(\leq h \text{ losses}) = [z^h] \frac{1}{1-z} \left(\begin{matrix} (1-c)z & cz \\ 1-b & b \end{matrix} \right)^n$$

with

$$c = (1-b) \frac{n-h}{n} .$$

Work in progress...

A Compound Poisson Model (1)

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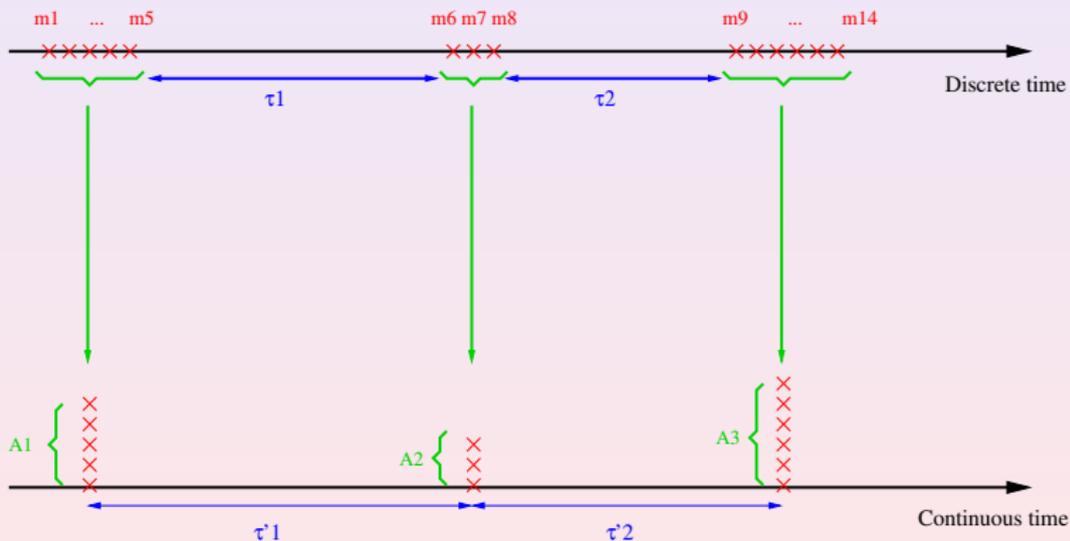
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Simplification: move to continuous time



A Compound Poisson Model (1)

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Process of loss:

- groups of losses occur according to a **Poisson process** with rate λ ,
- groups have random sizes with identical distribution and mean a .

Global loss rate: $p = \lambda \times a$

Distribution of the number of losses:

$$\sum_k z^k P(k \text{ losses in } [0, t)) = e^{\lambda(A(z)-1)} .$$

Comparison experiments (1)

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Comparison of two cases:

- Small bursts case: losses of
1 with proba 0.9,
2 with proba 0.1
- Larger bursts case: losses of
1 with proba 0.6,
2 with proba 0.4
- Same average packet loss number $x = p \times T$

$$\Delta_h(x) = P(\text{block saved with small bursts}) \\ - P(\text{block saved with larger bursts})$$

Comparison experiments (2)

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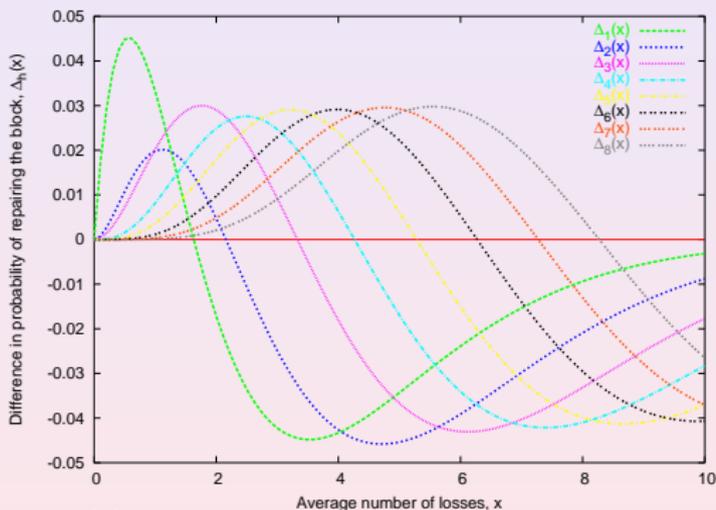
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Difference $\Delta_h(x)$ as the average number of losses x grows



Again an empirical law

$$x_h \sim h + C .$$

Analysis of limits

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Analysis of extreme cases: consider the probability of success of a block

$$P(N_T \leq h) = \sum_{n=0}^h \frac{x^n}{(\mathbb{E}A)^n n!} e^{-xT/\mathbb{E}A} P(A_1 + \dots + A_n \leq h).$$

i/ Assume that

$$\frac{\mathbb{P}(A^{(1)} > h)}{\mathbb{E}A^{(1)}} < \frac{\mathbb{P}(A^{(2)} > h)}{\mathbb{E}A^{(2)}}.$$

Then $\Delta_h(x) > 0$ when $x \rightarrow 0$.

ii/ Assume that $m^{(1)} < m^{(2)}$. Then $\Delta_h(x) < 0$ when $x \rightarrow \infty$.

Asymptotic Analysis (1)

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Consider the quantity:

$$d_h(y) = P(\leq h \text{ losses in } h + y \text{ time units}) .$$

Then we find:

$$d_h(y) = \frac{1}{2} + \frac{1}{\sqrt{2\pi h}} \sqrt{\frac{\mu_1}{\mu_2}} \left(\frac{1}{2} + \frac{\mu_3}{2\mu_2} - y \right) + o(h^{-1/2}) ,$$

where $\mu_1 = EA$, $\mu_2 = E(A^2)$, $\mu_3 = E(A^3)$.

Asymptotic Analysis (1)

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where $\mu_1 = EA$, $\mu_2 = E(A^2)$, $\mu_3 = E(A^3)$.

Asymptotic analysis (2)

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Accordingly: for all real y , we have:

$$\Delta_h(h+y) = \frac{1}{\sqrt{2\pi h}} (C_0 - C_1 y) + o(h^{-1/2}),$$

where:

$$C_0 = \sqrt{\frac{\mu_1^{(1)}}{\mu_2^{(1)}}} \left(\frac{1}{2} + \frac{\mu_3^{(1)}}{6\mu_2^{(1)}} \right) - \sqrt{\frac{\mu_1^{(2)}}{\mu_2^{(2)}}} \left(\frac{1}{2} + \frac{\mu_3^{(2)}}{6\mu_2^{(2)}} \right)$$

$$C_1 = \sqrt{\frac{\mu_1^{(1)}}{\mu_2^{(1)}}} - \sqrt{\frac{\mu_1^{(2)}}{\mu_2^{(2)}}}.$$

Asymptotic analysis (3)

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Finally, we have indeed:

$$\Delta_h(h + y_h) = 0 \implies y_h \sim \frac{C_1}{C_0},$$

and therefore

$$x_h \sim \frac{C_1}{C_0} + h.$$

FEC and the Queue Management

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As a conclusion

Packet queues inside network routers are handled by a **Queue Management** scheme.

Two common ones:

Tail Drop Drops packets if and only if the buffer is full
 \implies tends to produce bursts of losses

RED Drops packets at random preventively
 \implies tends to produce isolated losses

Two loss patterns: which one works better with FEC?

Application of the model

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Admitting that the smaller bursts (RED) work better when

$$x \leq h + C$$

for some constant C .

Equivalently, RED better if:

small block	$k \leq \frac{1-p}{p} h + \frac{C}{p}$
large redund. ratio	$\frac{h}{k} \geq \frac{1-p}{1-p} - \frac{C}{1-p} \frac{1}{k}$
small loss rate	$p \leq \frac{h+C}{h+k}$

Application of the model

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$$p(k+h) \leq h + C$$

for some constant C .

Equivalently, RED better if:

small block

$$k \leq \frac{1-p}{p} h + \frac{C}{p}$$

large redund. ratio

$$\frac{h}{k} \geq \frac{p}{1-p} - \frac{C}{1-p} \frac{1}{k}$$

small loss rate

$$p \leq \frac{h+C}{h+k} .$$

Experimental setup

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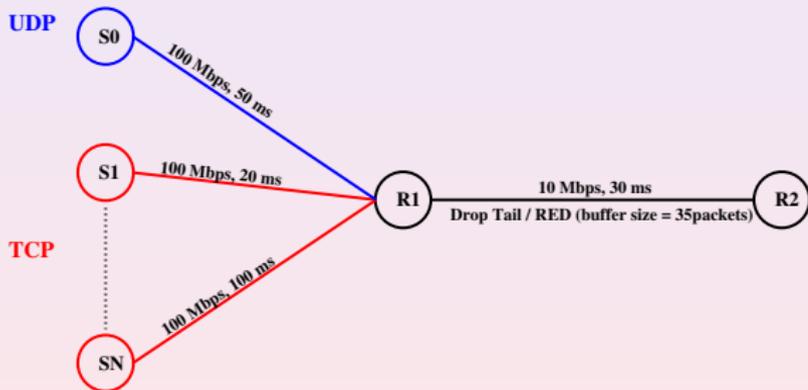
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Simulations with the ns-2 program.

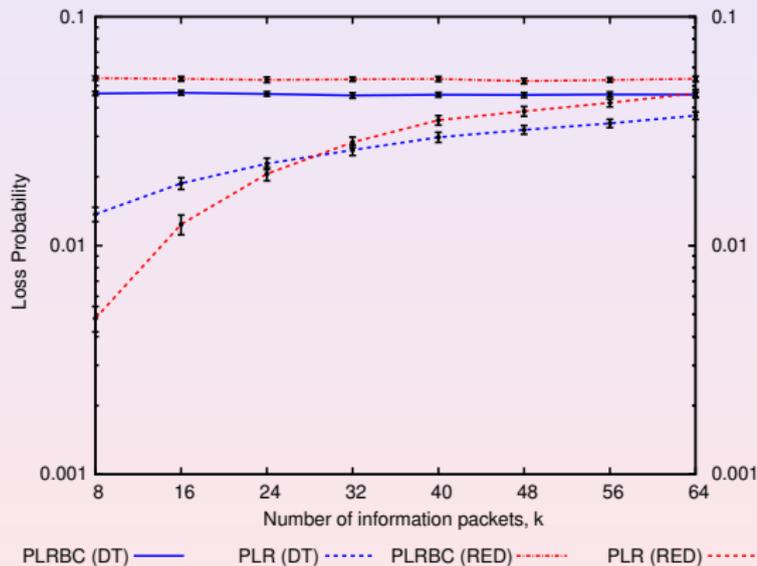
- Source of packets with the UDP protocol, 5-10% of the BW
- Background traffic of TCP flows, saturating the BW.



Statistics collected about Packet Loss Rate **Before Correction** and **after correction**.

Results of Simulations

Loss rates, $k = 16$ packets per block + $h = 2$ FEC packets.



RED does not always win...

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There is a compromise between loss “burstiness” and loss rate.
Assume blocks protected with $h = 1$ packet.

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