

The Laser model: structure and properties

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Advanced Markov Modeling
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 - Stationary distribution
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Constituents:

- $2n$ energy levels, forming two “bands”; at most one electron occupies an energy level;
- two special “lasing levels” L et ℓ in the middle of both bands
- a cavity: supports photons.

Transitions (changes of state):

- electrons may change level in each band (thermalization)
- one electron may emit one photon, and passes simultaneously from L to ℓ
- one electron may absorb a photon, and passes simultaneously from ℓ à L
- photons may exit the cavity
- “pumping” makes an electron pass from the lowest to the highest energy level.

What is a state of the system? At least two choices:

electron-centric : each electron has a state, its energy level $\in [0..2n - 1]$. The number of photons is *a priori* unbounded.

$$\implies \mathcal{E} \subset [0..2n - 1]^N \times \mathbb{N}.$$

But the exclusion constraint has to be enforced.

level-centric : each energy level contains at most one electron.

$$\implies \mathcal{E} \subset \{0, 1\}^{2n} \times \mathbb{N}.$$

Since the total number of electrons is fixed:

$$\implies \mathcal{E} \subset \left\{ \sigma \in \{0, 1\}^{2n} \mid \sum_{i=0}^{2n-1} \sigma_i = N \right\} \times \mathbb{N}.$$

The time scales of thermalization and other events (emission, absorption, light, pumping) are very different.

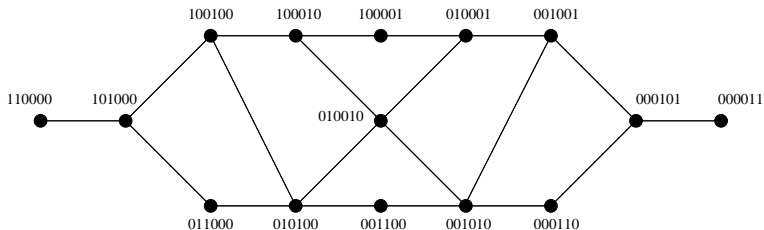
The process may be seen as:

- 2 independent processes in each of the bands
- coupled by *rare* events

⇒ focus on each band

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Example of a transition diagram: $n = 6$ levels, $N = 2$ particles.



All transitions represented are two-way. Unwritten transition rates: rate p for transitions to the left (lower energy), pq to the right (higher energy).

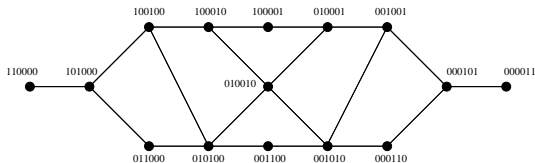
Order states lexicographically: matrices can be represented as

$$M_{n,m} = \left(\begin{array}{c|cc} & \mathbf{0}_{p \times q} & \mathbf{0}_{p \times r} \\ \hline & \lambda & \\ & \ddots & \\ & & \lambda & \mathbf{0}_{q \times r} \\ \hline M_{n-1,m-1} & & & \\ \hline \mathbf{0}_{q \times p} & \mu & & \\ & & \ddots & \\ & & & \mu & \\ \hline \mathbf{0}_{r \times p} & \mathbf{0}_{r \times q} & & \\ \hline & & & M_{n-1,m} \end{array} \right) .$$

The base matrix is the $n \times n$ generator of the *birth and death* process with n states: (for the laser, $\lambda = p$ and $\mu = pq$)

$$M_{n,1} = \begin{pmatrix} - & \lambda & & & \\ \mu & - & \lambda & & \\ & \ddots & \ddots & \ddots & \\ & & \mu & - & \lambda \\ & & & \mu & - \end{pmatrix}.$$

But if we order by energy level, we get a “quasi-birth-death” process

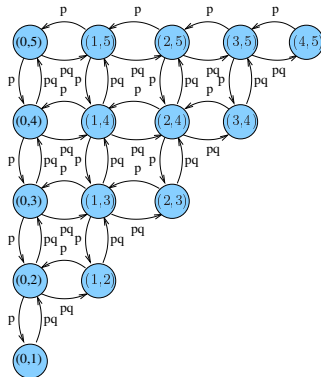


The matrix is then: $M_{6,2} =$

-	λ																		
μ	-	λ	λ																
	μ	-	0		λ														
	μ	0	-		λ														
		μ	0		-	0		λ	λ	0									
		μ	μ		0	-		0	λ	λ									
			μ	0	μ	0		0	0	0		λ	0						
				μ	μ	0		0	0	-		0	0		λ	0			
					μ	μ	0		0	-		0	-		λ	λ			
						0	μ	μ			0	-		μ	0		λ		
											μ	μ			0	-	λ		
														μ	μ	-		λ	
																μ	-		

And what happens if we adopt the electron-centric model?

$$\mathcal{E} = \left\{ \mathbf{X} \in [0..2n-1]^N \mid X_1 < X_2 < \dots < X_N \right\}$$



For instance, still with $n = 6$, $P = 2$, a regular, 2-dimensional grid-like structure

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The stationary distribution can be computed in *closed form* as:

$$\pi_{n,m}(\boldsymbol{\sigma}) = \frac{1}{G(n,m)} \prod_{i=0}^{n-1} q^{i\sigma_i} = \frac{1}{G(n,m)} q^{\sum_{i=0}^{n-1} i\sigma_i} = \frac{q^{U\boldsymbol{\sigma}}}{G(n,m)}.$$

The *partition function* $G(n, m)$ is given by:

$$\begin{aligned} G(n, m) &= \sum_{\boldsymbol{\sigma} \in \mathcal{S}_{n,m}} \prod_{i=0}^{n-1} q^{i\sigma_i} \\ &= \frac{q^{m-1} - q^n}{1 - q^m} \frac{q^{m-2} - q^n}{1 - q^{m-1}} \cdots \frac{1 - q^n}{1 - q}. \end{aligned}$$

Proofs:

- direct check of balance equations
- reversibility and truncation

Definition (Reversible Markov Chains)

A continuous-time Markov chain is reversible if $\{X(t)\}$ have $\{X(-t)\}$ the same generator, that is, if for all pair of states $(i, j) \in \mathcal{E} \times \mathcal{E}$,

$$\pi(i) q_{i,j} = \pi(j) q_{j,i} .$$

This equation is the **local balance** condition. It implies that when transition $j \rightarrow k$ exists, transition $k \rightarrow j$ must exist as well. It is sufficient that it be satisfied for *some* family of numbers that need not be, *a priori* a probability distribution. If it does, then this family turns out to be indeed a stationary distribution for \mathbf{Q} , possibly up to a normalization constant.

The class of Reversible Markov chains has property very useful for computing distributions: it is stable with respect to *truncation*.

Definition (Truncated Markov Chain)

Consider a CTMC over a countable state space \mathcal{E} and generator \mathbf{Q} . Let $\mathcal{A} \subset \mathcal{E}$. The truncated Markov chain over the sub-space \mathcal{A} is the CTMC with generator $\mathbf{Q}_{\mathcal{A}}$ defined as:

$$(q_{\mathcal{A}})_{jk} = \begin{cases} q_{jk} & \text{if } j, k \in \mathcal{A} \\ 0 & \text{otherwise.} \end{cases} \quad (1)$$

In words, the transitions of the process are the same as the original process as long as the state is in \mathcal{A} . But transitions outside \mathcal{A} have rate 0. If the process starts inside \mathcal{A} , it stays there forever, and it can be considered as a CTMC on the space \mathcal{A} .

Theorem (Stationary distribution of truncated reversible chains)

Let $\mathcal{A} \subset \mathcal{E}$. Consider some reversible CTMC, with stationary distribution π . Then the truncated Markov chain over \mathcal{A} is also reversible, and its stationary distribution is given by: for $i \in \mathcal{A}$

$$\pi'(i) = \frac{\pi(i)}{\sum_{j \in \mathcal{A}} \pi(j)} .$$

Open question: is it possible to calculate marginal probabilities (occupancies). For each energy level i ,

$$\mathbb{P}\{N_i = 1\} = ?$$

The direct approach is:

$$\begin{aligned} \mathbb{P}\{N_i = 1\} &= \sum_{\sigma \in \mathcal{S}_{n,m} \mid \sigma_i=1} \pi_{n,m}(\sigma) \\ &= \frac{1}{G(n,m)} \sum_{\sigma \in \mathcal{S}_{n,m} \mid \sigma_i=1} q^{\sum_{j=0}^{n-1} j \sigma_j} . \end{aligned}$$

Too complex \implies need for some simpler computation, e.g. with a recurrence.

Two recurrences

$$\mathbb{P}\{N_i = 1, N + 1 \text{ particles}\} = q^i \frac{1 - q^{N+1}}{q^N - q^n} (1 - \mathbb{P}\{N_i = 1, N \text{ particles}\}).$$

$$\begin{aligned} \mathbb{P}\{N_i = 1\} \\ &= \frac{\rho^i}{G(n, N)} \sum_{k=0}^i \rho^{(i+1)(N-1-k)} G(i, k) G(n - i - 1, N - k - 1) \end{aligned}$$

$$\begin{aligned} \mathbb{P}\{N_i = 0\} \\ &= \frac{1}{G(n, N)} \sum_{k=0}^i \rho^{(i+1)(N-k)} G(i, k) G(n - i - 1, N - k). \end{aligned}$$

Sampling from the stationary distribution

It is possible to *directly* sample from this stationary distribution, although the state space is very large.

This is because the distribution has a *recursive* structure.

The “simplex” $\mathcal{S}_{n,m}$ can be decomposed as:

$$\mathcal{S}_{n,m} = \{1\} \times \mathcal{S}_{n-1,m-1} \cup \{0\} \times \mathcal{S}_{n-1,m}, \quad 0 < m < n.$$

Then we get the recursion:

$$\begin{aligned} G(n, m) &= q^{m-1} G(n-1, m-1) + q^m G(n-1, m). \\ \pi_{n,m}(\sigma_0 = 1) &= q^{m-1} \frac{G(n-1, m-1)}{G(n, m)} \\ \pi_{n,m}(\sigma_0 = 0) &= q^m \frac{G(n-1, m)}{G(n, m)} \end{aligned}$$

+ boundary values.

Sampling from the stationary distribution, forward method

Data: Two integers n and m , $0 \leq n \leq m$

Result: A vector in $\mathcal{S}_{n,m}$, sampled from the distribution $\pi_{n,m}$

begin

$k \leftarrow m$

for i **from** 0 **to** $n - 1$ **do**

if $k = 0$ **then**

$\sigma[i] \leftarrow 0$

else if $k = n - i$ **then**

$\sigma[i] \leftarrow 1$

else

$u \leftarrow \text{Uniform}([0, 1])$

if $u < q^k \frac{G(n-1-i, k)}{G(n-i, k)}$ **then**

$\sigma[i] \leftarrow 0$

else

$\sigma[i] \leftarrow 1$

$k \leftarrow k - 1$

return $(\sigma[0], \dots, \sigma[n-1])$

Symmetrically, $\mathcal{S}_{n,m}$ is decomposed as

$$\mathcal{S}_{n,m} = \mathcal{S}_{n-1,m-1} \times \{1\} \cup \mathcal{S}_{n-1,m} \times \{0\}, \quad 0 < m < n.$$

Then we get:

$$\begin{aligned} G(n, m) &= q^{n-1} G(n-1, m-1) + G(n-1, m) \\ \pi_{n,m}(\sigma_{n-1} = 1) &= q^{n-1} \frac{G(n-1, m-1)}{G(n, m)} \\ \pi_{n,m}(\sigma_0 = 0) &= \frac{G(n-1, m)}{G(n, m)}. \end{aligned}$$

+ boundary values.

Sampling from the stationary distribution, backwards method

Data: Two integers n and m , $0 \leq n \leq m$

Result: A vector in $\mathcal{S}_{n,m}$, sampled from the distribution $\pi_{n,m}$

begin

$k \leftarrow m$

for i **from** 1 **to** n **do**

if $k = 0$ **then**

$\sigma[n-i] \leftarrow 0$

else if $k = n - i$ **then**

$\sigma[n-i] \leftarrow 1$

else

$u \leftarrow \text{Uniform}([0, 1])$

if $u < \frac{G(n-i, k)}{G(n-i+1, k)}$ **then**

$\sigma[n-i] \leftarrow 0$

else

$\sigma[n-i] \leftarrow 1$

$k \leftarrow k - 1$

return $(\sigma[0], \dots, \sigma[n-1])$

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The question is to quantify the distance between $\pi_t = \pi_0 e^{\mathbf{Q}t}$ and the stationary distribution π .

The *spectral decomposition* principle says that if $\mathbf{Q} = \mathbf{R}\mathbf{D}\mathbf{R}^{-1}$, then

$$\begin{aligned} e^{\mathbf{Q}t} &= \mathbf{R}e^{\mathbf{D}t}\mathbf{R}^{-1} = \mathbf{R} \begin{pmatrix} 1 & & & 0 \\ & e^{\lambda_2 t} & & \\ & & \ddots & \\ 0 & & & e^{\lambda_n t} \end{pmatrix} \mathbf{R}^{-1} \\ &= \mathbf{1}\pi + \sum_{j=2}^n e^{\lambda_j t} \mathbf{A}_j . \end{aligned}$$

Locating the eigenvalues (spectrum) λ_j gives information on the speed of convergence.

The basis matrix is:

$$M_{n,1} = \begin{pmatrix} -\lambda & \lambda & & & & & \\ \mu & -(\lambda + \mu) & \lambda & & & & \\ & \ddots & \ddots & \ddots & & & \\ & & & \mu & -(\lambda + \mu) & \lambda & \\ & & & & \mu & -\mu & \end{pmatrix}.$$

Its n eigenvalues are 0 and:

$$\omega_k = -(\lambda + \mu) + 2\sqrt{\lambda\mu} \cos \frac{k\pi}{n}, \quad k = 1..n - 1.$$

Speed of convergence: given by the largest, non-zero one. It is less than

$$-(\lambda + \mu) + 2\sqrt{\lambda\mu} = -(\sqrt{\lambda} - \sqrt{\mu})^2.$$